

# Introduction to Local Estimation

# Overview

1. Traditional vs. piecewise SEM
2. Tests of directed separation
3. Log-likelihood assessment
4. Introduction to *piecewiseSEM*
5. A Warning

# 1.1 Traditional vs. Piecewise SEM

# 1.1 Comparison. Traditional vs. piecewise SEM

Variance-covariance	Piecewise
Single (global) variance-covariance matrix estimated	Variance-covariance matrices estimated for each endogenous variable
Simultaneous solution (computationally intensive)	Multiple solutions (modularized)
Fit to normal distribution	Incorporates various distributions (Poisson, Gamma, etc.)
Assumes independence	Can model non-independence (blocked, temporal, spatial, etc.)
Latent & composite variables	No latent variables (yet*)
Recursive & non-recursive (cyclic)	Only for recursive (acyclic)
Multi-group models	Can estimate random components, but no formal $\chi^2$ test

# 1.1 Comparison. Traditional vs. piecewise SEM

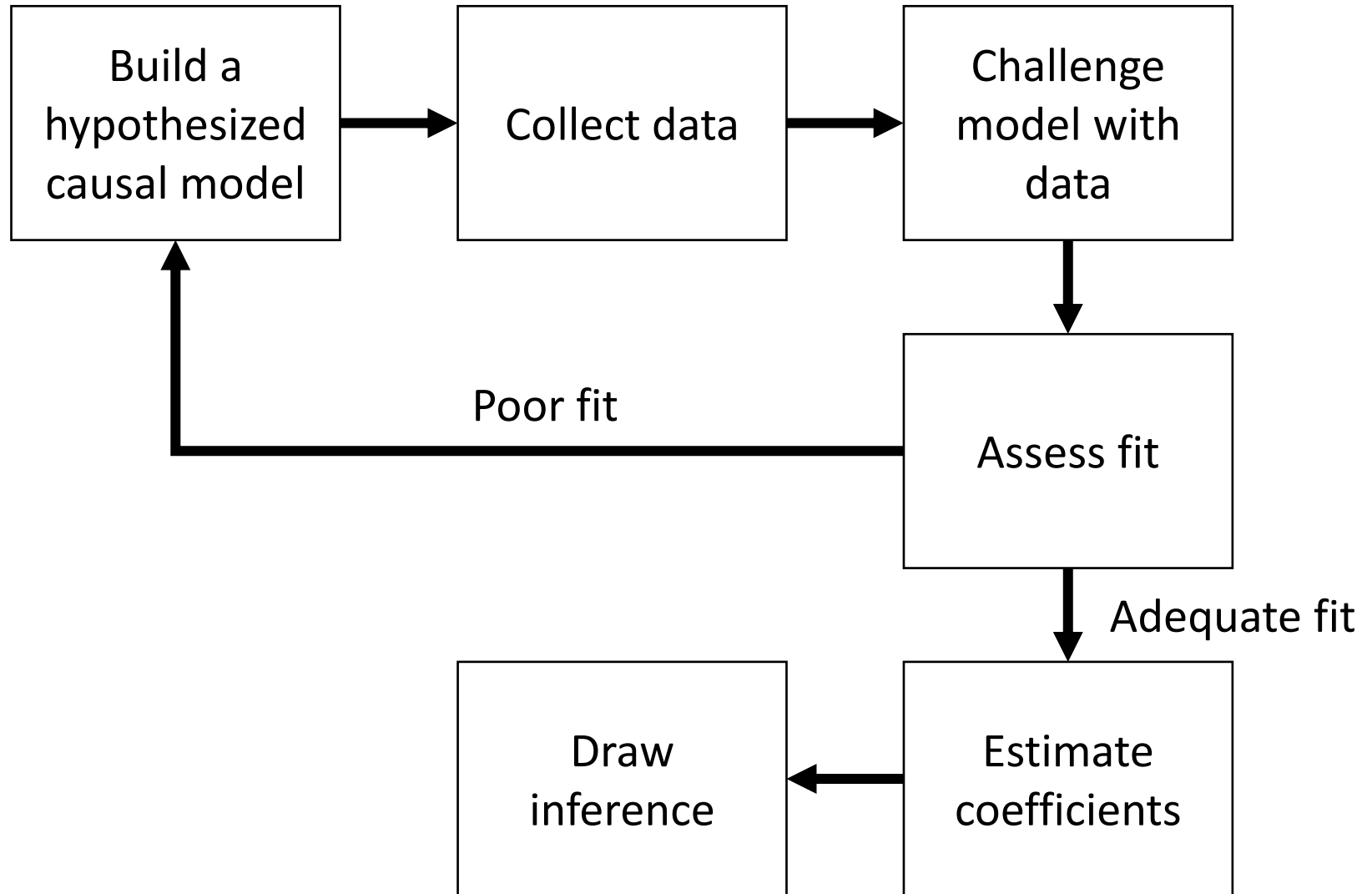
## Traditional SEM



## Piecewise SEM



# 1.1 Piecewise SEM.



# 1.2 Tests of Directed Separation

## 1.2 Directed Separation. Model fit

Does the model fit the data?

=

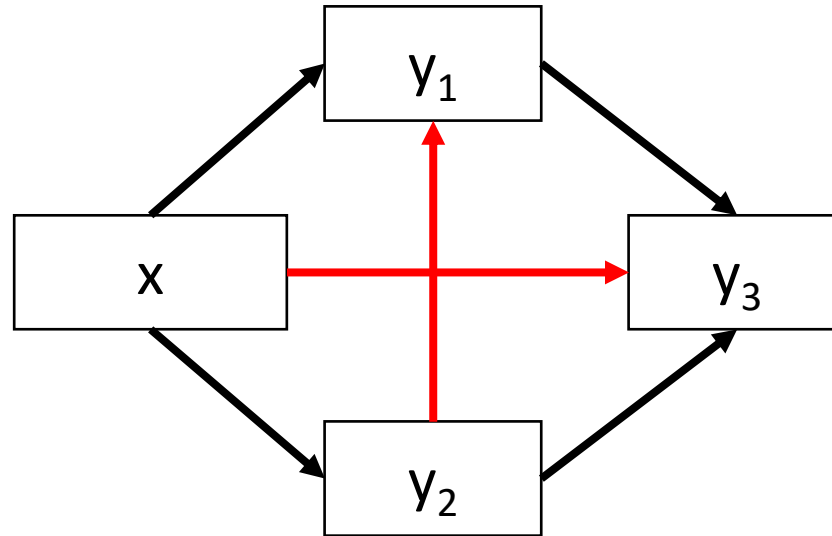
Does the model represent the data well?

=

Are we missing important information?

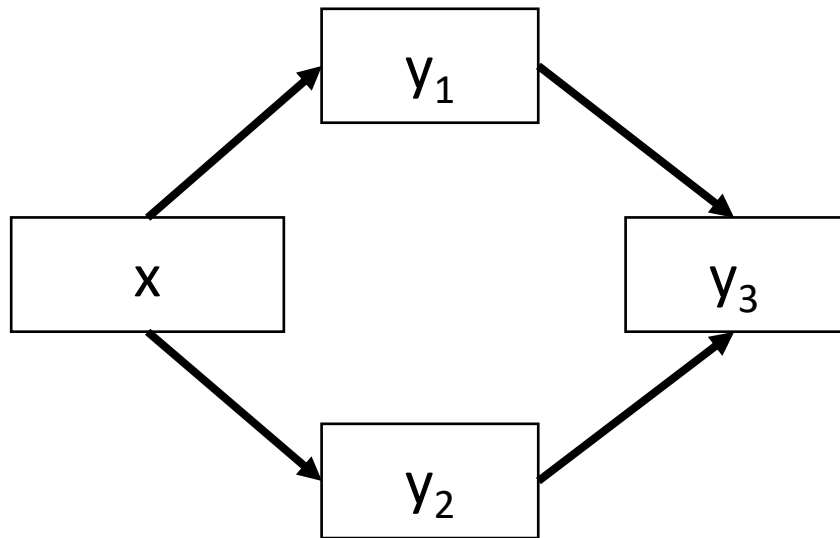


## 1.2 Directed Separation. Model fit



Did we get the topology right or are there unrecognized significant relationships?

## 1.2 Directed Separation. D-separation



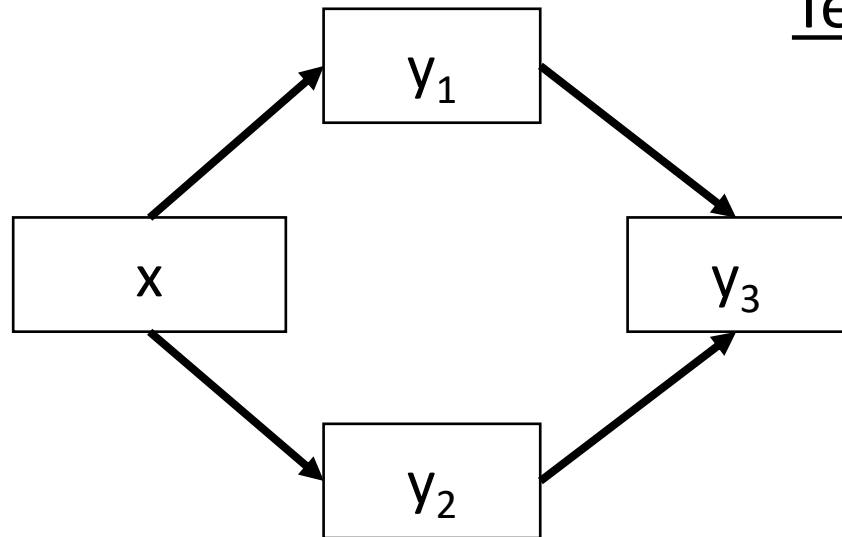
- Concept from *Graph Theory*
- Two nodes are d-separated if they are *conditionally independent* e.g., the effect of  $x$  on  $y_3$  is zero conditioning on the influences of  $y_1$  and  $y_2$

## 1.2 Directed Separation. Independence claims

The d-separation criterion for any pair of variables involves:

1. *Directly* controlling for causal connections via *immediate* parents
2. *Indirectly* controlling for common ancestors that could generate correlations between the pair (as reflected in immediate parents)
3. Not controlling for common descendent variables (the effect of the cause)

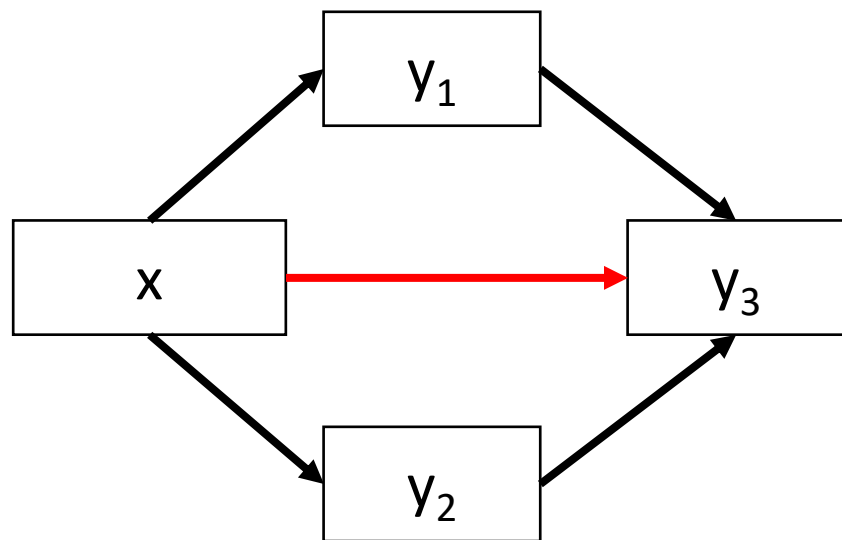
## 1.2 Directed Separation. D-separation



Test of directed separation:

1. Identify all *conditional independence* claims
2. Evaluate each independence claim
3. Summarize information across all claims

## 1.2 Directed Separation. Independence claims



### 1. Identify all independence claims

*Basis set* = the *smallest* possible set of independence claims from a graph

*Notation:*

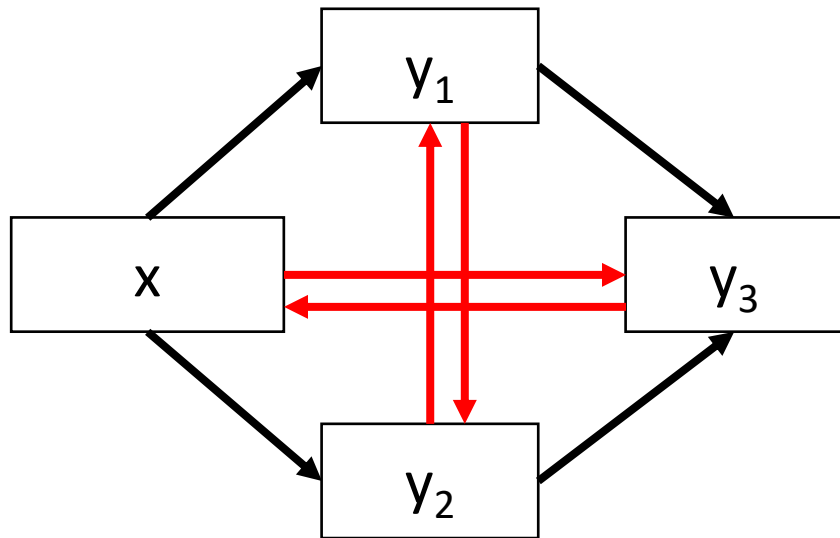
$$x \perp y_3 \mid (y_1, y_2)$$

Predictor involved in claim

Response

Conditioning variables

## 1.2 Directed Separation. Independence claims



### 1. Identify all independence claims

$$1. x \perp y_3 \mid (y_1, y_2)$$

~~$$2. y_3 \perp x \mid (y_1, y_2)$$~~

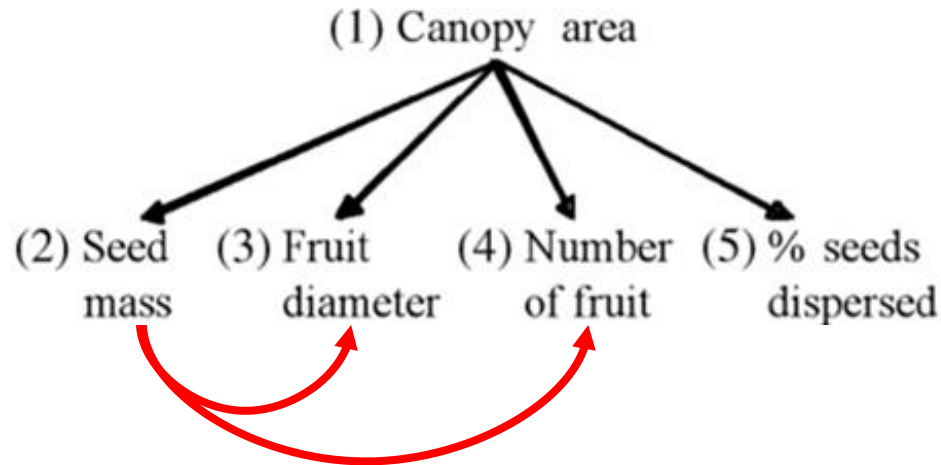
$$3. y_1 \perp y_2 \mid (x)$$

~~$$4. y_2 \perp y_1 \mid (x)$$~~

“Minimum set” = reciprocal relationships are not part of the basis set ( $P$ -values are identical in either direction, EXCEPT...)

## 1.2 Directed Separation. Deriving the basis set

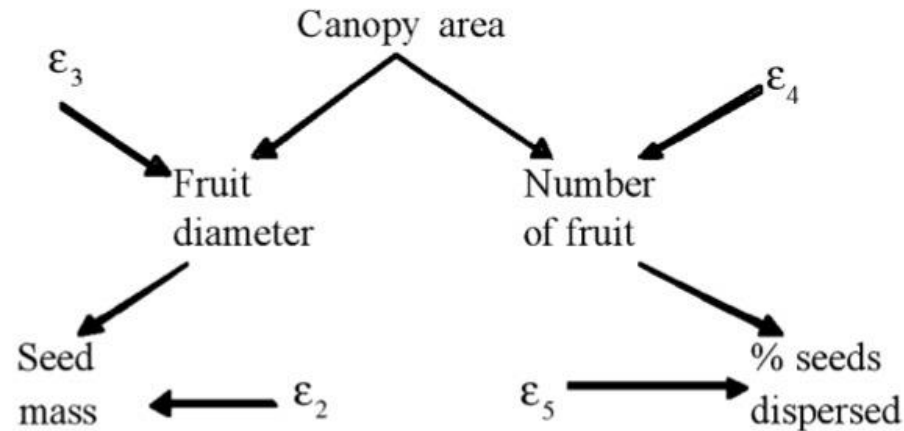
What is the basis set?



- |  |   |
|--|---|
| 1. $\text{mass} \perp \text{dia} \mid (\text{canopy})$ | 4. $\text{dia} \perp \# \mid (\text{canopy})$ |
| 2. $\text{mass} \perp \# \mid (\text{canopy})$         | 5. $\text{dia} \perp \% \mid (\text{canopy})$ |
| 3. $\text{mass} \perp \% \mid (\text{canopy})$         | 6. $\% \perp \# \mid (\text{canopy})$         |

## 1.2 Directed Separation. Deriving the basis set

What is the basis set?

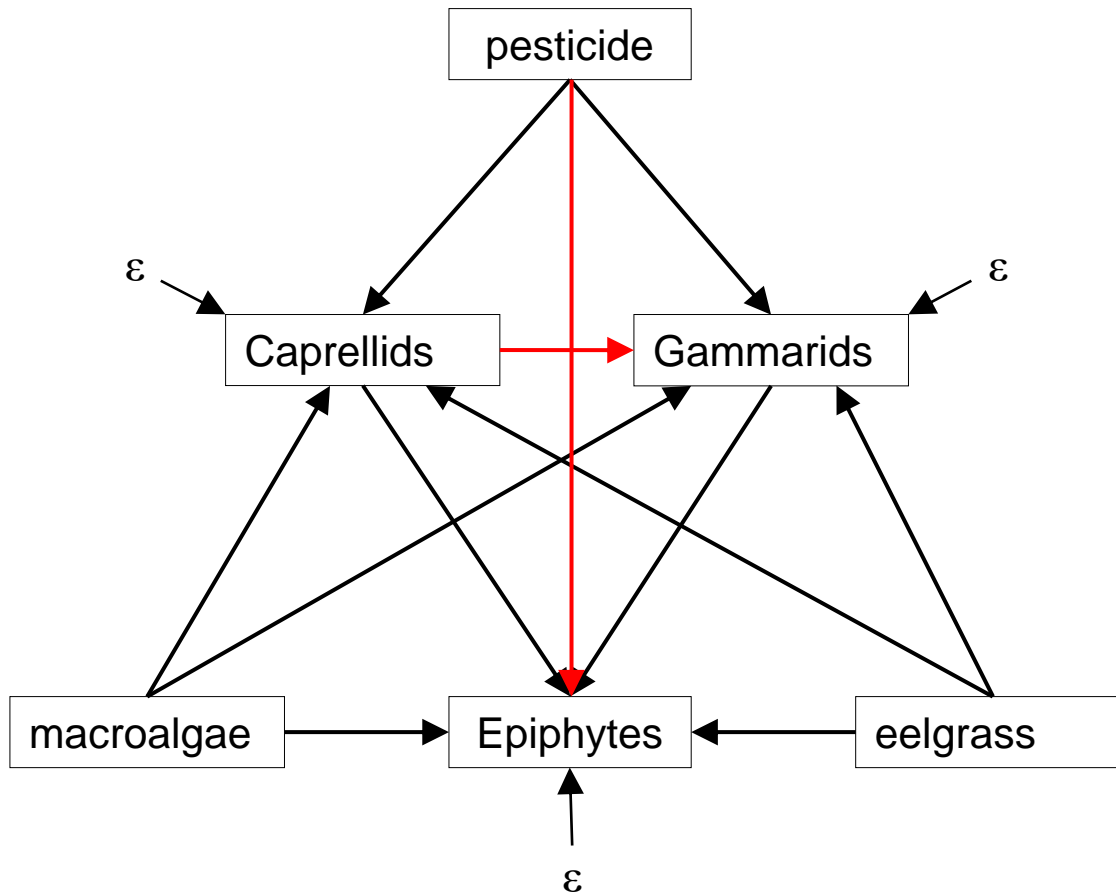


1.  $\text{dia} \perp \# \mid (\text{canopy})$
2.  $\text{dia} \perp \% \mid (\text{canopy}, \#)$
3.  $\text{canopy} \perp \text{mass} \mid (\text{dia})$
4.  $\text{mass} \perp \# \mid (\text{dia}, \text{canopy})$
5.  $\text{mass} \perp \% \mid (\text{dia}, \#)$
6.  $\text{canopy} \perp \% \mid (\#)$



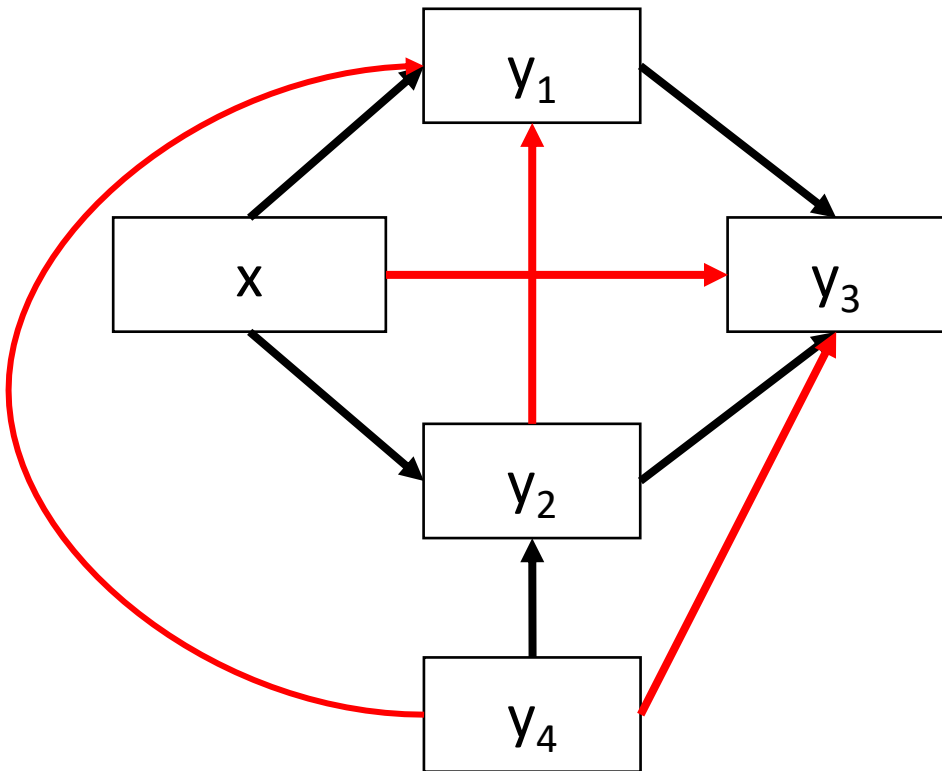
# 1.2 Directed Separation. Deriving the basis set

## What is the basis set?



1. Pesticide  $\perp$  epiphytes  
(macroalgae, eelgrass, caprellids, gammarids)
2. Caprellids  $\perp$  gammarids  
(macroalage, eelgrass, pesticide)

## 1.2 Directed Separation. Independence claims

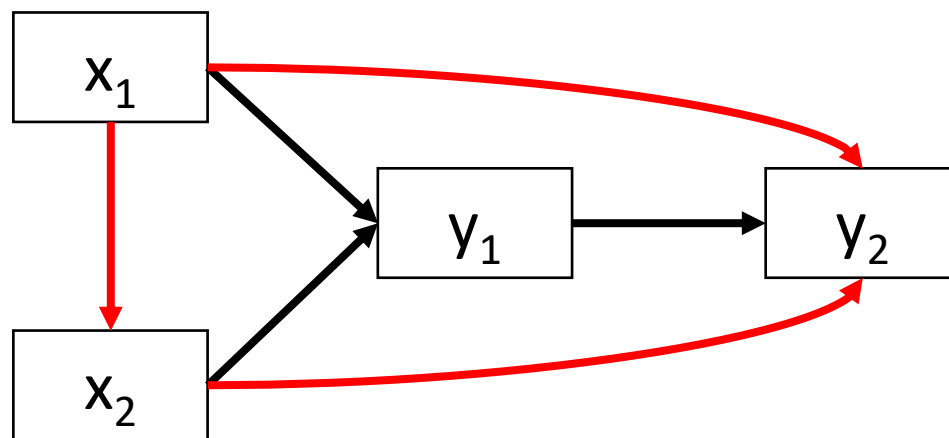


### 1. Identify all independence claims

1.  $x \perp y_3 \mid (y_1, y_2)$  ↖
2.  $y_1 \perp y_2 \mid (x)$  Not conditioning on  $y_4$ , since its >1 node away
3.  $y_4 \perp y_1 \mid (x)$
4.  $y_4 \perp y_3 \mid (y_1, y_2)$
5.  $y_4 \perp x???$

## 1.2 Directed Separation. A note

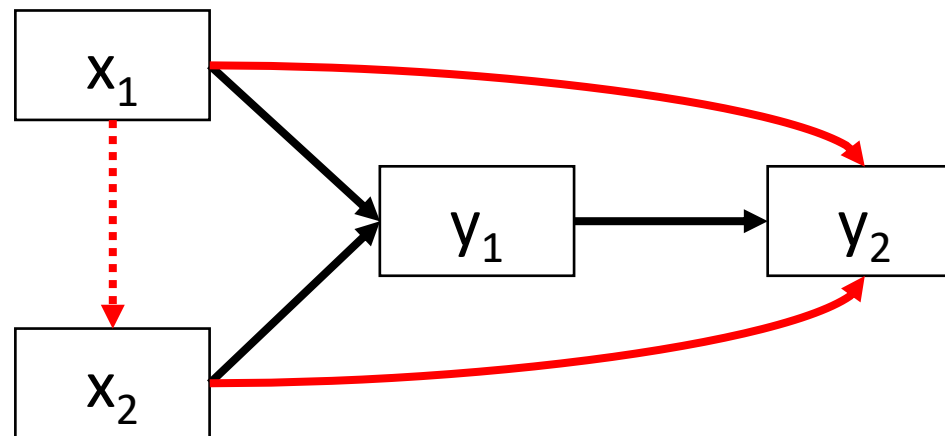
- Basis set excludes relationships among exogenous variables



1.  $x_1 \perp y_2 \mid (y_1)$
2.  $x_2 \perp y_2 \mid (y_1)$

## 1.2 Directed Separation. A note

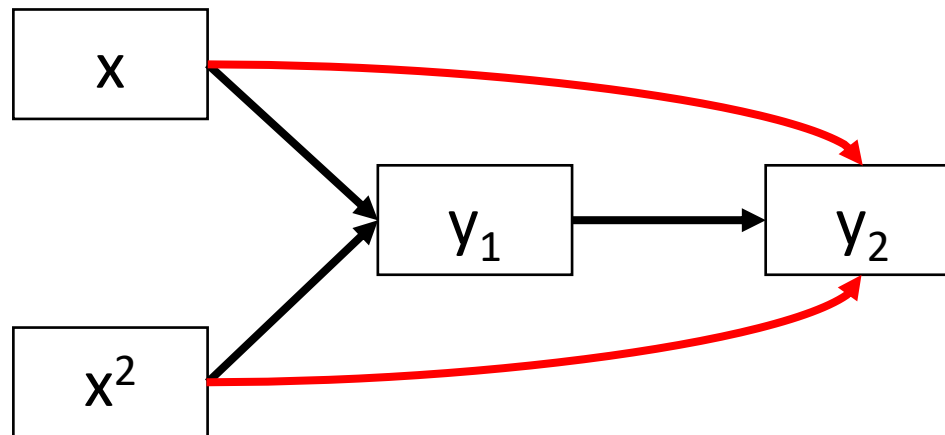
- Unclear as to the direction of the relationship ( $x_1 \rightarrow x_2$  or  $x_2 \rightarrow x_1$ )
- Unclear whether variables could even be plausibly causally linked (e.g., ocean basin and latitude)
- Distributional assumptions, etc. not defined



1.  $x_1 \perp y_2 \mid (y_1)$
2.  $x_2 \perp y_2 \mid (y_1)$

## 1.2 Directed Separation. A note

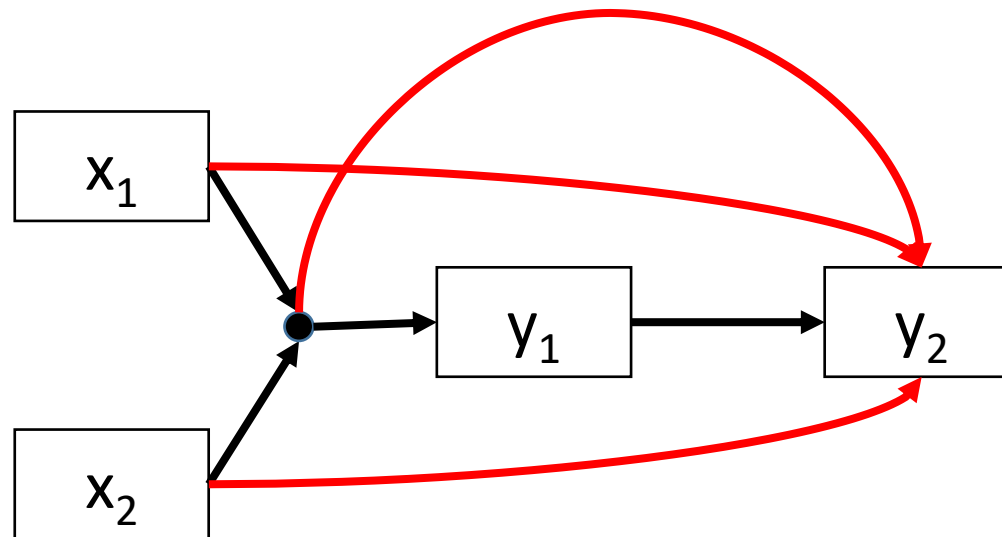
- Basis set generally excludes non-linear components (polynomials)



1.  $x \perp y_2 \mid (y_1)$

## 1.2 Directed Separation. A note

- Basis set generally excludes non-linear components (interactions)

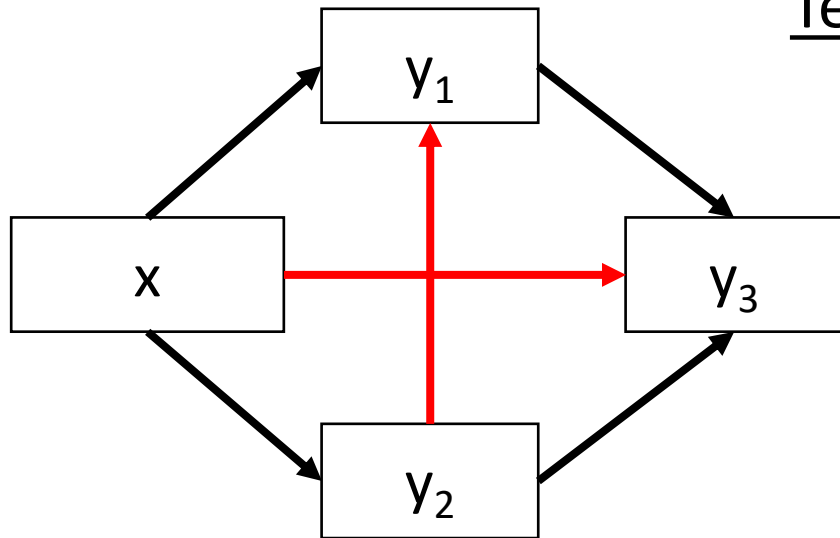


1.  $x_1 \perp y_2 \mid (y_1)$
2.  $x_2 \perp y_2 \mid (y_1)$

# ACTIVITY

- Take your causal diagram from Day 1
- Derive the basis set

## 1.2 Directed Separation. D-separation

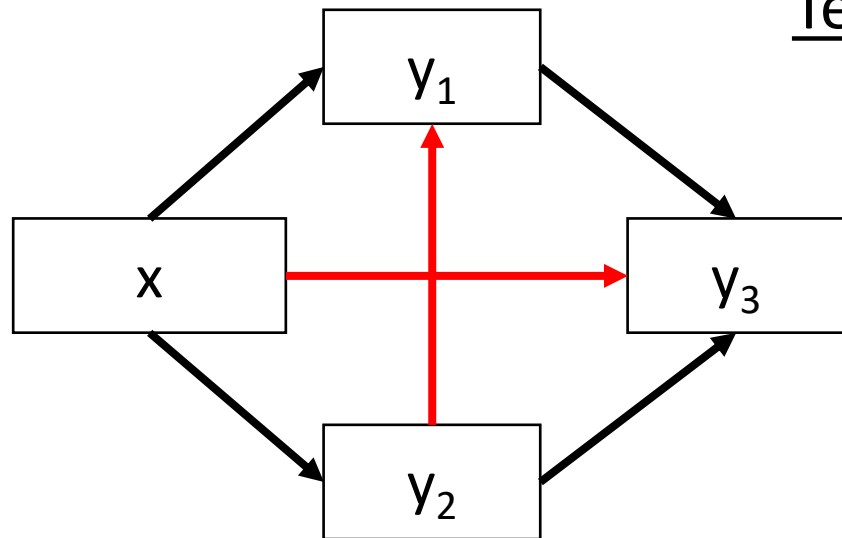


Test of directed separation:

1. Identify all independence claims
2. **Evaluate each independence claim**
3. Summarize information across all claims



## 1.2 Directed Separation. D-separation



Test of directed separation:

$$1. x \perp y_3 \mid (y_1, y_2)$$

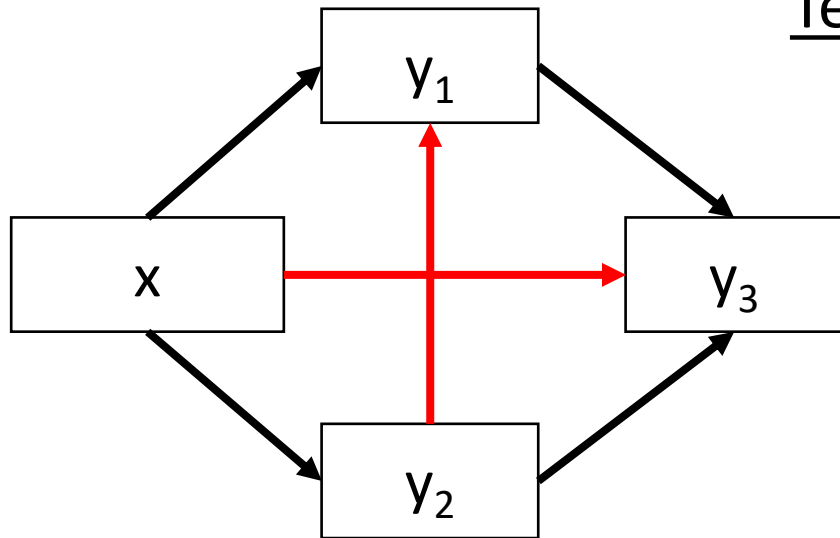
$$y_3 \sim y_1 + y_2 + x$$

$$1. y_1 \perp y_2 \mid (x)$$

$$y_2 \sim x + y_1$$

- Fit models (using same parameters as originally specified) and extract null significance statistic:  $P$ -value)
- A non-significant  $P$ -value suggests the claim is conditionally independent (i.e., relationship is no different than 0)

## 1.2 Directed Separation. D-separation



### Test of directed separation:

1. Identify all independence claims
2. Evaluate each independence claim
3. **Summarize information across all claims**

## 1.2 Directed Separation. Fisher's $C$

- Summarize independence claims across basis set:

$$C = -2 * \sum \ln(p_i)$$

$p_i$  = the  $P$ -values of all tests of conditional independence

- $C$  has a  $\chi^2$ -square distribution with  $2k$  degrees of freedom
- $k$  = # of elements of the basis set

## 1.2 Directed Separation. Fisher's $C$

What is  $p < 0.05$ ?

- You are likely missing some associations
- You *reject* this model
- The way forward: adding links or different model structure? (look at d-sep tests)
- To re-iterate,  $p \geq 0.05$  is GOOD

## 1.2 Directed Separation. Model selection

- Fisher's  $C$  can be used to construct model AIC:

$$AIC = C + 2K$$

- $K$  = # of likelihood parameters estimated (not to be confused with  $k$ )
- Can be extended to small sample size:

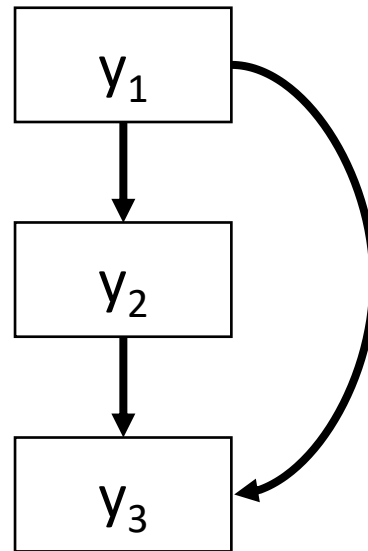
$$AIC_c = C + 2K(n / (n - K - 1))$$

## 1.2 Directed Separation. Complexity and sample size

- Shipley suggests need only enough individual degrees of freedom to fit each component model
- Or,  $d$ -rule (Grace et al 2015):
  - $d = \# \text{ of samples} / \# \text{ of pathways}$
  - $d \geq 5$
- More is always better...
  - Low sample size leads to non-significant  $d$ -sep tests
  - Low sample size also leads to non-significant path coefficients
  - End up with a 'good fitting' model that says *nothing*

## 1.2 Directed Separation. Saturated models

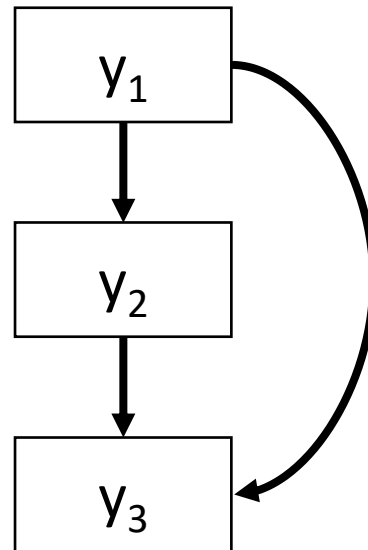
What is the basis set?



- There is no basis set for a saturated model (all paths are represented)
- No d-sep tests and therefore no  $C$  statistic can be constructed for this model (same as global estimation)

## 1.2 Directed Separation. Saturated models

What is the basis set?



Options:

1. Remove the mediating variable  $y_2$  and test submodel
2. Rely on other indicators of model fit (e.g., path significance,  $R^2$ )

**A global test is not the be all-end all of models**



## 1.3 Log-likelihood assessment

## 1.3. Model fit

Does the model fit the data?

=

Does the model represent the data well?

=

Are there more *likely* configurations?

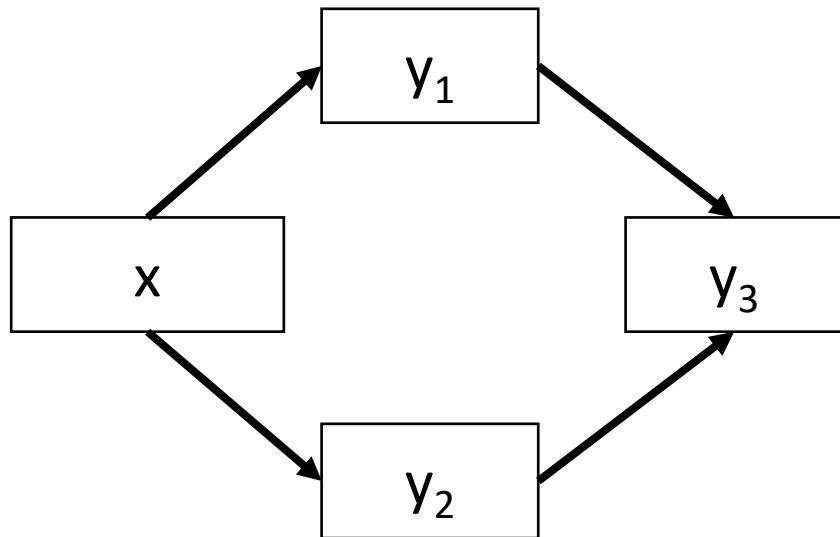
## 1.3. Directed separation. The trouble with $P$ -values

- $P$ -values are not always returned by default (see: *lme4*)
- D-sep tests only reflect changes in *topology* (whether paths or variables are missing), but there are lots of components to tweak (distributions, transformations, link functions, etc.) that don't affect the topology

## 1.3. Log-likelihood

- *Recall*: Maximum likelihood estimation = find the parameters (coefficients) that maximize the probability of observing the data
- Likelihood = value of the maximum likelihood fitting function with the optimal parameters
- Log-likelihood (L-L) = the log-transformation of the likelihood

## 1.3. Log-likelihood



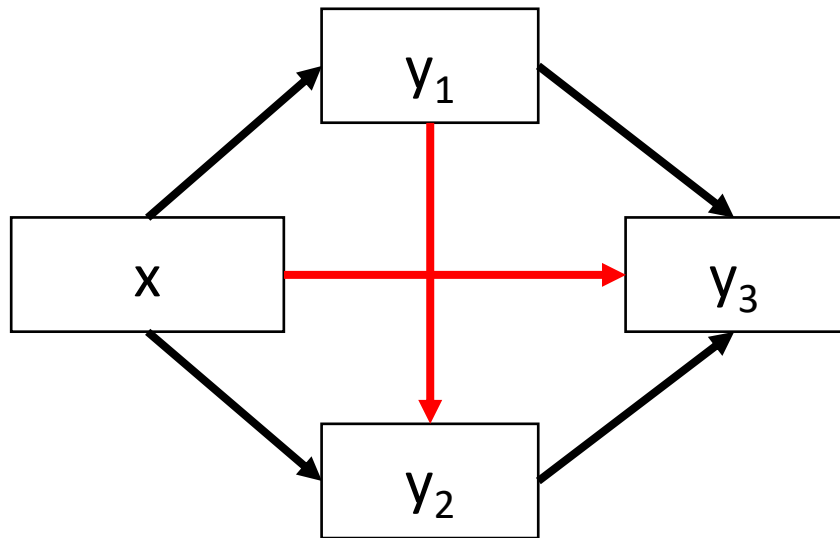
Can compute L-L for each model in our SEM:

$$y_1 \sim \gamma_1 x$$

$$y_2 \sim \gamma_2 x$$

$$y_3 \sim \beta_1 y_1 + \beta_2 y_2$$

## 1.3. Log-likelihood



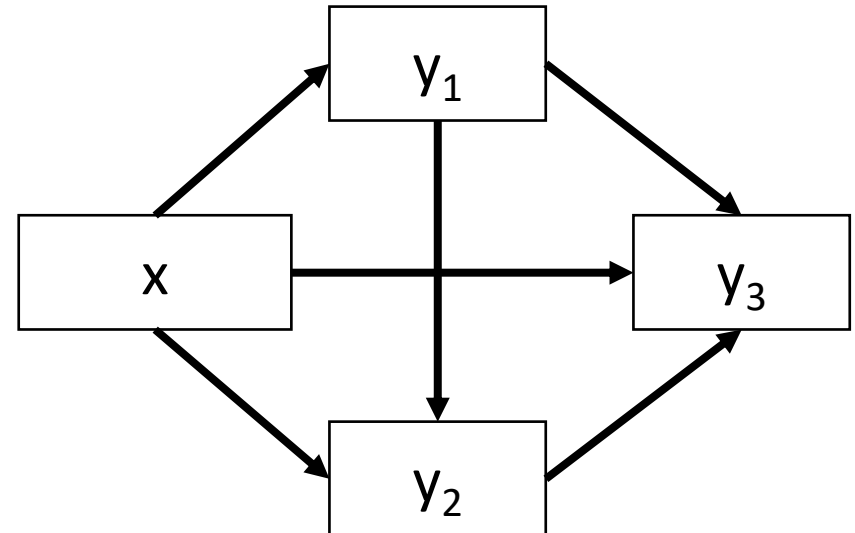
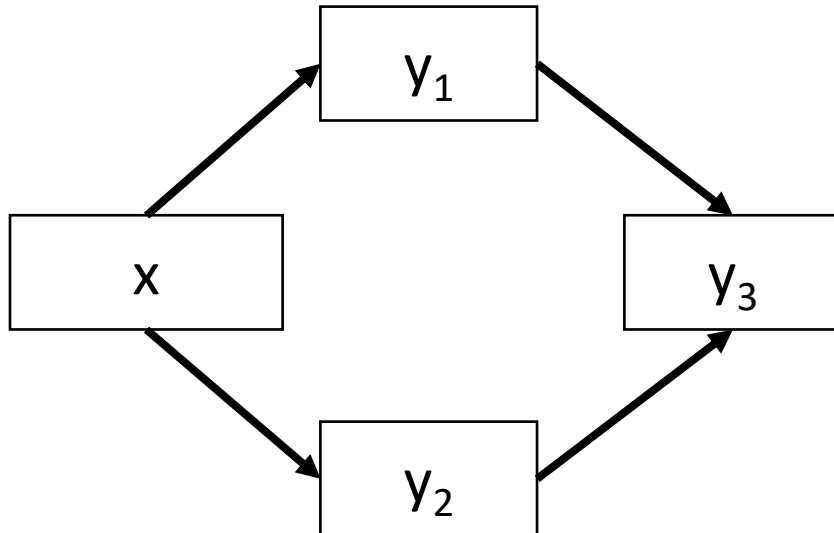
- Recall the goal of d-sep tests are to evaluate missing paths = same as asking if the paths are *no different than 0*
- What is the alternate hypothesis? That these paths *are* different from zero

# 1.3. Log-likelihood. Competing models

Proposed causal model

VS

Saturated model

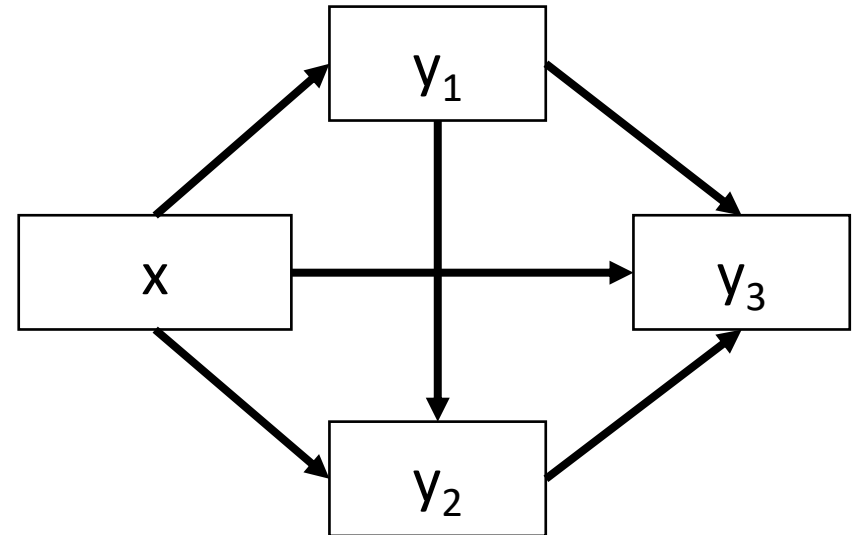
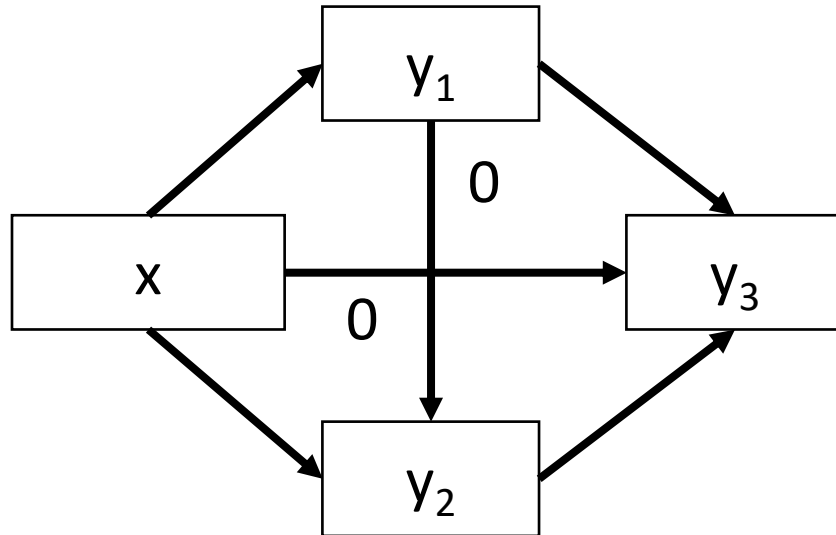


# 1.3. Log-likelihood. Competing models

Proposed causal model

VS

Saturated model



(this is also what we test in global estimation... is the difference in the estimated and observed covariances zero?)

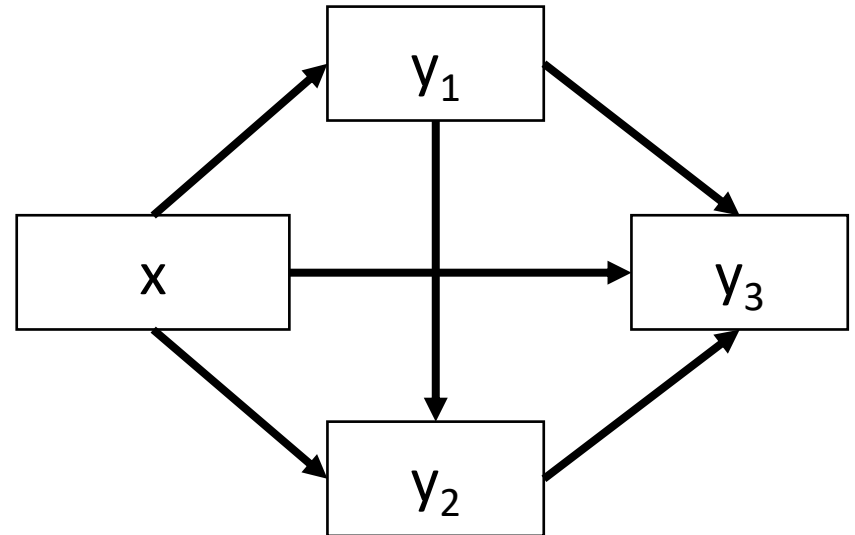
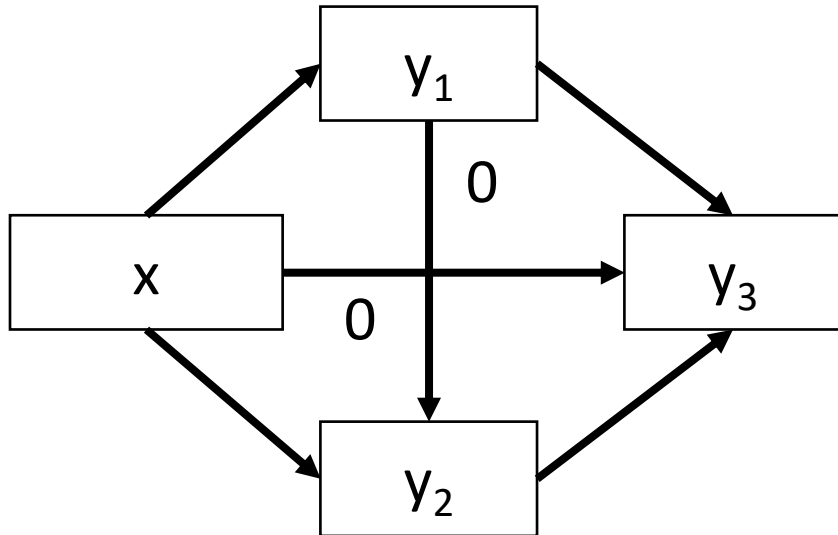


# 1.3. Log-likelihood. Competing models

Proposed causal model

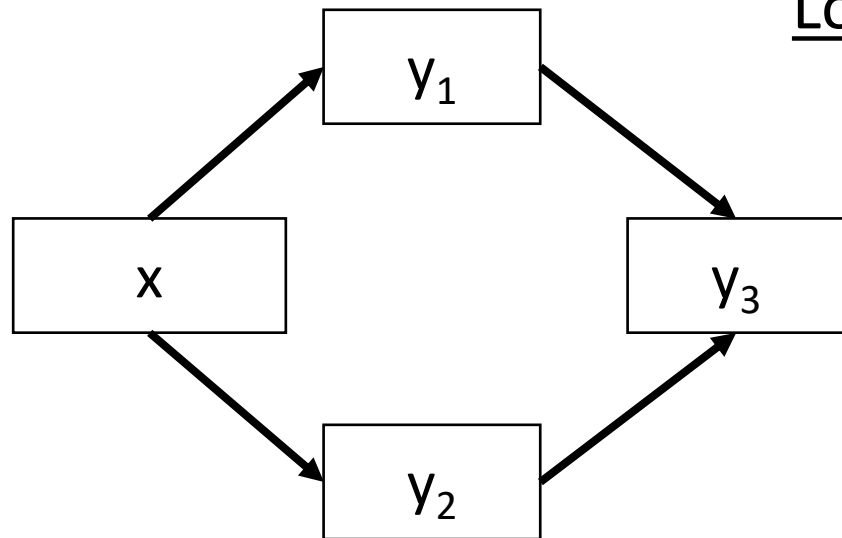
VS

Saturated model



How much more **likely** is the model in which all paths are free to vary (saturated model) than the proposed model?

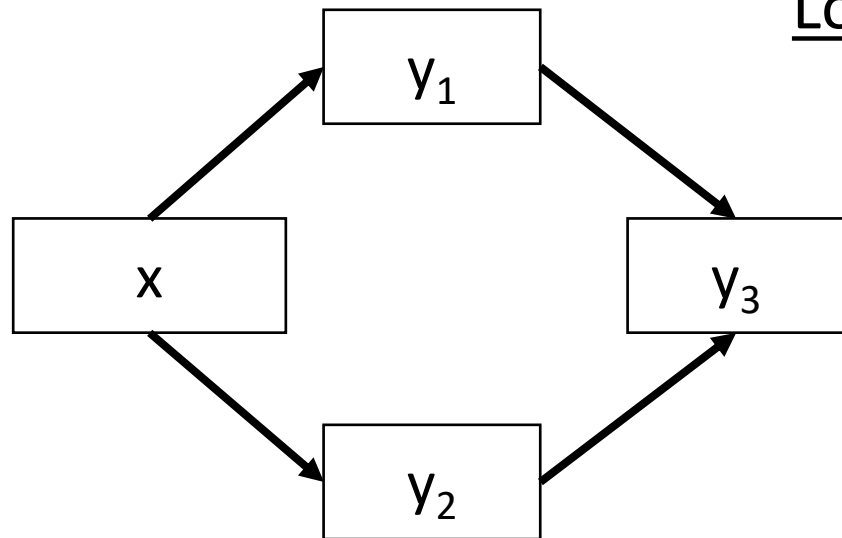
## 1.3. Log-likelihood. LRT



### Log-likelihood test:

1. Fit proposed model and summarize log-likelihoods
2. Fit saturated model and summarize log-likelihoods
3. Test whether they are different

## 1.3. Log-likelihood. LRT



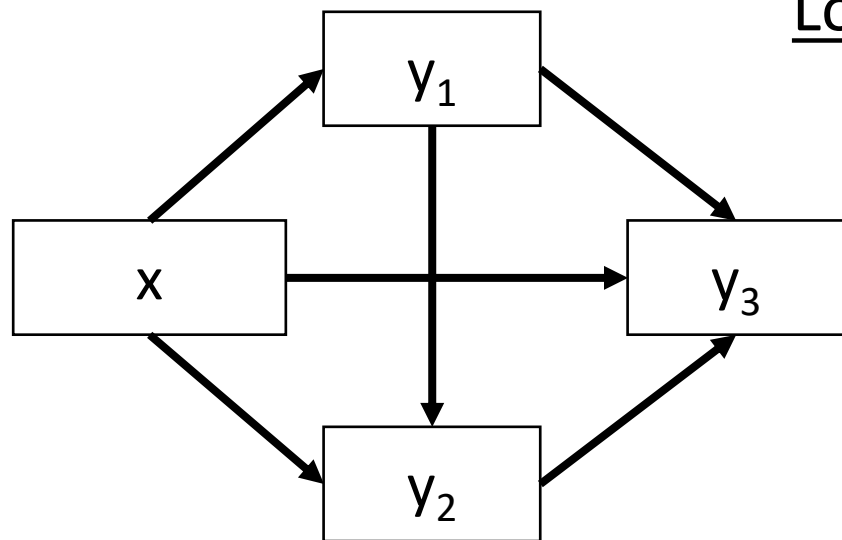
Log-likelihood test:

- 1. Fit proposed model and summarize log-likelihoods**

Log-likelihood of SEM is the sum of the individual log-likelihoods

$$\log \mathcal{L}_M(\theta_M|X) = \sum_{i=1}^v \log(\mathcal{L}_i(\theta_i|X))$$

## 1.3. Log-likelihood. LRT

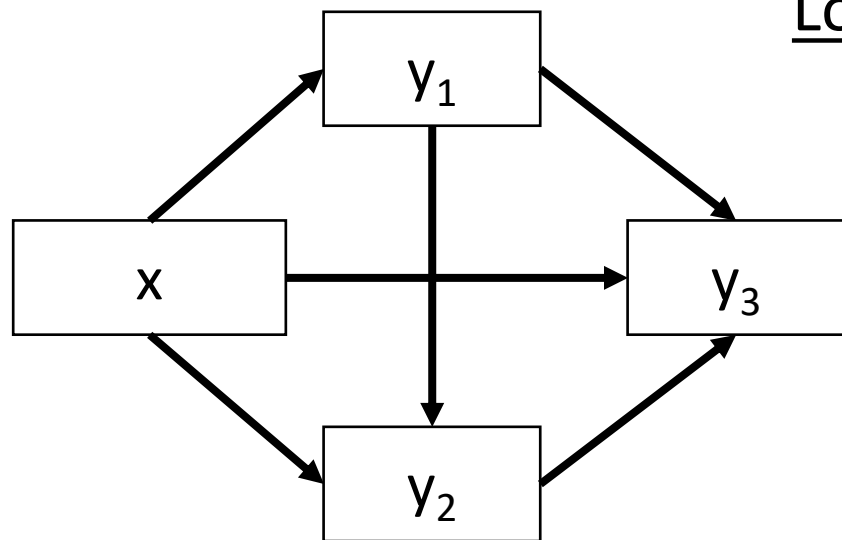


Log-likelihood test:

1. Fit proposed model and summarize log-likelihoods
- 2. Fit saturated model and summarize log-likelihoods**
3. Test whether they are different

$$\log \mathcal{L}_M(\theta_M|X) = \sum_{i=1}^v \log(\mathcal{L}_i(\theta_i|X))$$

## 1.3. Log-likelihood. LRT



### Log-likelihood test:

1. Fit proposed model and summarize log-likelihoods
2. Fit saturated model and summarize log-likelihoods
3. **Test whether they are different** =  $\chi^2$  likelihood ratio test

$$\chi^2 = -2(\log(\mathcal{L}(M_1)) - \log(\mathcal{L}(M_2)))$$

## 1.3. Log-likelihood. Goodness-of-fit

- $\chi^2$  statistic is the *same as we get from global estimation if* we assume multivariate normality
- LRT requires that proposed model be nested within the saturated model
- Allows extensions of techniques from global estimation (e.g., modification indices = how much does  $\chi^2$  likelihood change with additional removal of paths?)
- Can be extended to any model that uses ML estimation (e.g., GAMs) that were previously prohibited
- *Cannot* be used with techniques that are not estimated using maximum likelihood (e.g., quasi-likelihood)

## 1.3. Log-likelihood. Model comparison

- Can extend likelihood summing concept to compute model-wide AIC from submodel AICs:

$$AIC_M = \sum_{i=1}^v AIC_i$$

- Can be extended to small sample size correction
- Solves issue with lack of AIC from d-sep tests for saturated models

## 1.3. Log-likelihood. Issues

- If model does not converge or random effects are close to 0, then can produce wonky log-likelihood estimates
- This can lead to the impossible situation where  $\chi^2 < 0$ 
  - In this case , you will get an NA for  $\chi^2$  statistic
- What to do?
  - Re-fit model and tweak optimization parameters to encourage convergence
  - Drop random effects whose variance components are very small from the model
  - Revert to d-sep tests



# 1.4 Introduction to *piecewiseSEM*



Local\_Estimation.R

## 1.4 piecewiseSEM.

***piecewiseSEM***: Piecewise structural equation modeling in R for ecology, evolution, and systematics

```
install.packages("devtools")  
library(devtools)  
install_github("js1efche/piecewiseSEM@devel")
```

*When you see this, time to code along!*

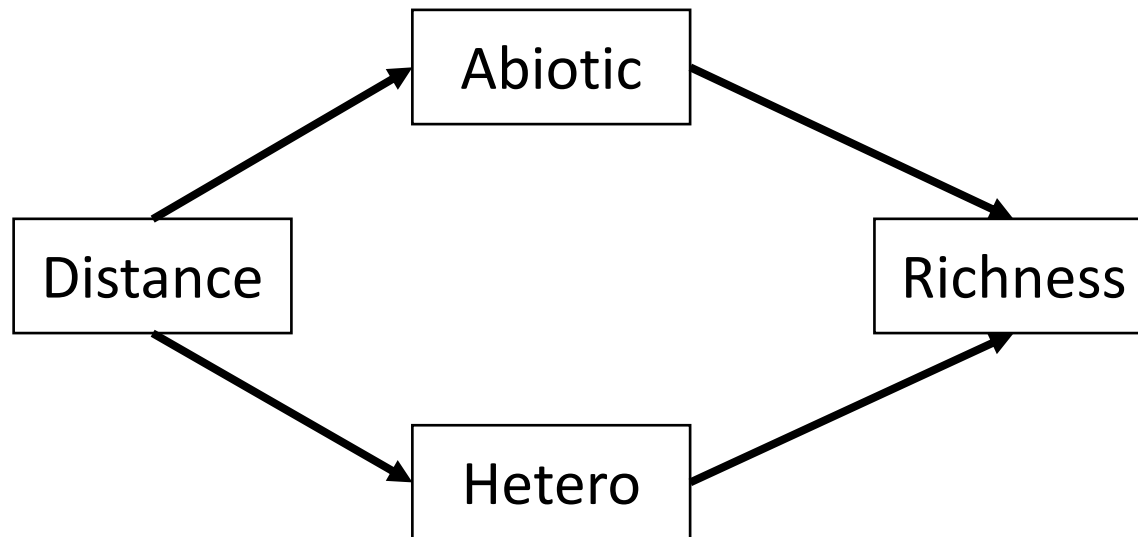


# Mediation in Analysis of Post-Fire Recovery of Plant Communities in California Shrublands



Five year study of wildfires in Southern California in 1993.  
90 plots (20 x 50m)

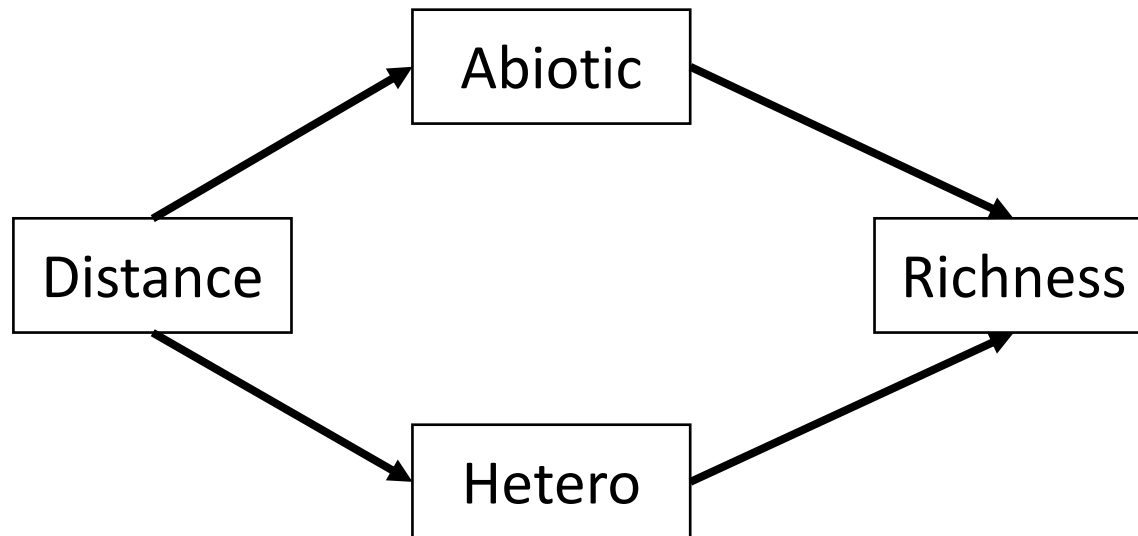
## 1.4 piecewiseSEM. Keeley example



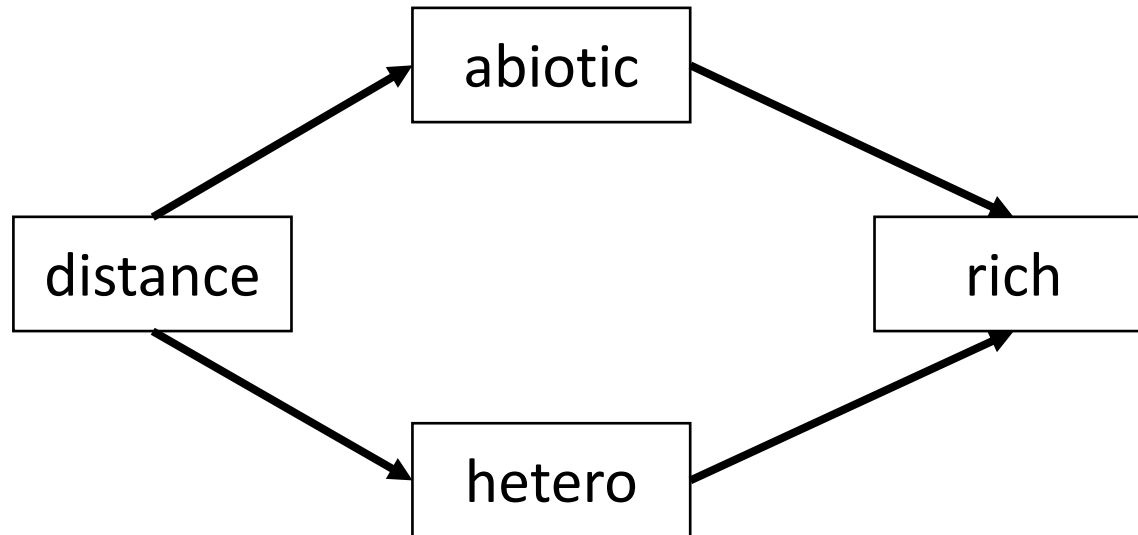
1. Create list of structured equations
2. Conduct d-sep tests (evaluate fit)
3. Construct  $\chi^2$  from log-likelihoods
4. Extract coefficients

## 1.4 piecewiseSEM. Keeley example

Break this model up into component models



## 1.4 piecewiseSEM. Store list of equations



```
# Read in data
data(keeley)

# Create list of structured equations
keeley.sem <- psem(
  lm(abiotic ~ distance, data = keeley),
  lm(hetero ~ distance, data = keeley),
  lm(rich ~ abiotic + hetero, data = keeley),
  data = keeley
)
```

# 1.4 piecewiseSEM. Store list of equations

```
keeley.sem
```

```
Structural Equations of x :
```

```
lm: abiotic ~ distance
```

```
lm: hetero ~ distance
```

```
lm: rich ~ abiotic + hetero
```

```
Data:
```

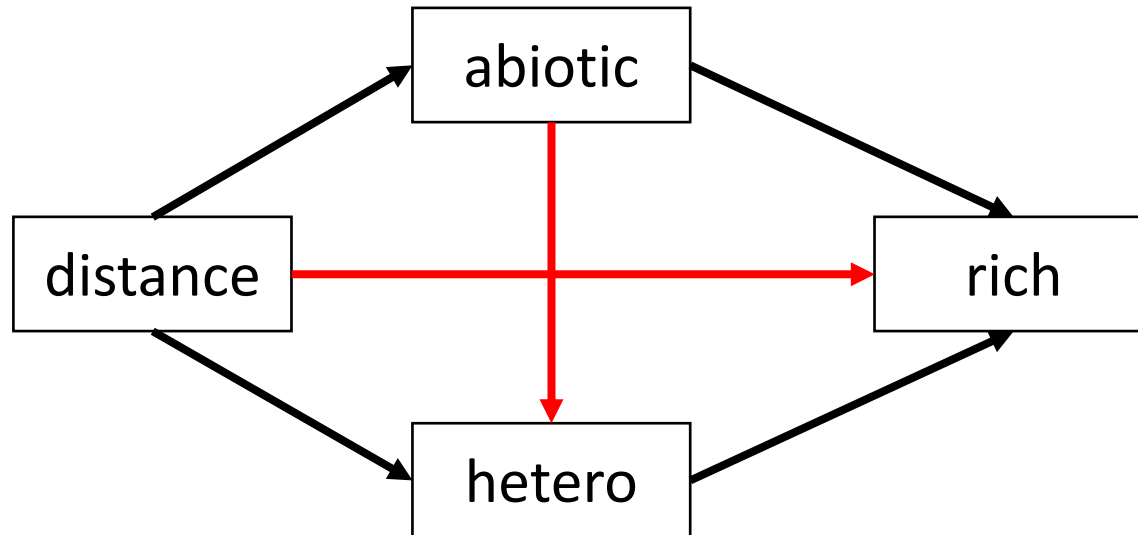
	distance	elev	abiotic	age	hetero	firesev	cover	rich
1	53.40900	1225	60.67103	40	0.757065	3.50	1.0387974	51
2	37.03745	60	40.94291	25	0.491340	4.05	0.4775924	31
3	53.69565	200	50.98805	15	0.844485	2.60	0.9489357	71
4	53.69565	200	61.15633	15	0.690847	2.90	1.1949002	64
5	51.95985	970	46.66807	23	0.545628	4.30	1.2981890	68
6	51.95985	970	39.82357	24	0.652895	4.00	1.1734866	34

...with 84 more rows

```
[1] "class(psem)"
```



## 1.4 piecewiseSEM. D-sep tests



```
# Get the basis set
basisSet(keeley.sem)

$`1`
[1] "distance | rich ( abiotic, hetero )"

$`2`
[1] "abiotic | hetero ( distance )"
```



## 1.4 piecewiseSEM. D-sep tests

```
# Conduct d-sep tests
```

```
claim1 <- lm(rich ~ distance + abiotic + hetero, keeley)
```

```
coefs(claim1)
```

	Response	Predictor	Estimate	Std.Error	DF	Crit.Value	P.Value	Std.Estimate	
1	rich	distance	0.6404	0.1565	86	1.0933	0.0001	0.3743	***
2	rich	abiotic	0.5233	0.1756	86	2.9793	0.0038	0.2660	**
3	rich	hetero	33.4010	11.1187	86	3.0040	0.0035	0.2539	**

```
claim2 <- lm(hetero ~ abiotic + distance, keeley)
```

```
coefs(claim2)
```

	Response	Predictor	Estimate	Std.Error	DF	Crit.Value	P.Value	Std.Estimate	
1	hetero	abiotic	0.0022	0.0017	87	1.4296	0.1871	0.1491	
2	hetero	distance	0.0036	0.0015	87	2.4742	0.0153	0.2774	*

```
# Compute Fisher's C & compare to Chi-square distribution
```

```
C <- -2 * (log(coefs(claim1)[1, 7]) + log(coefs(claim2)[1, 7]))
```

```
1 - pchisq(C, 2 * 2)
```

```
[1] 0.0002223955
```



## 1.4 piecewiseSEM. D-sep tests

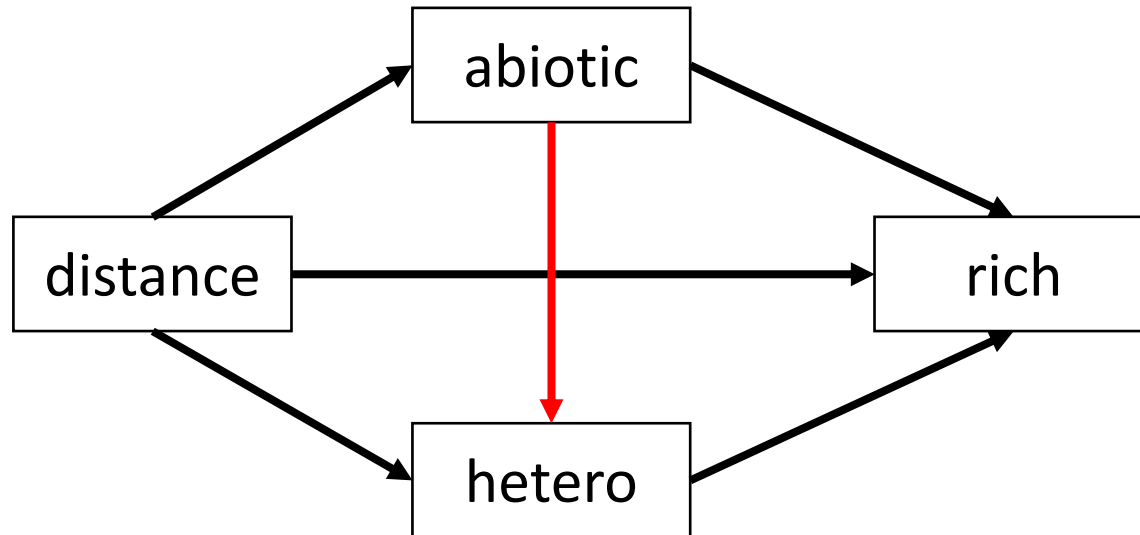
```
# Magically conduct d-sep tests  
(keeley.dsep <- dSep(keeley.sem))
```

```
# By default the conditioning variables are hidden, but we can show  
them  
dSep(keeley.sem, conditioning = TRUE)
```

Independ.Claim	Test.Type	DF	Crit.Value	P.Value
1 rich ~ distance + abiotic + hetero	coef	86	4.093329	9.564005e-05 ***
2 hetero ~ abiotic + distance	coef	87	1.329585	1.871306e-01



## 1.4 piecewiseSEM. Re-assess fit

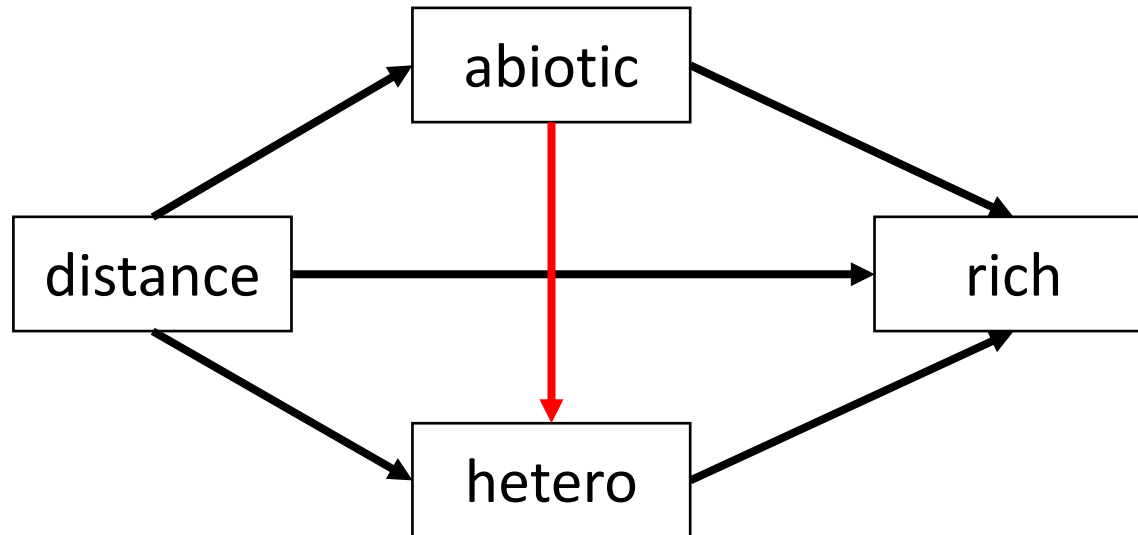


```
# Add significant path back into model
keeley.sem2 <- update(keeley.sem, rich ~ abiotic + hetero + distance)

dSep(keeley.sem2)

fisherC(keeley.sem2)
```

## 1.4 piecewiseSEM. Re-assess fit



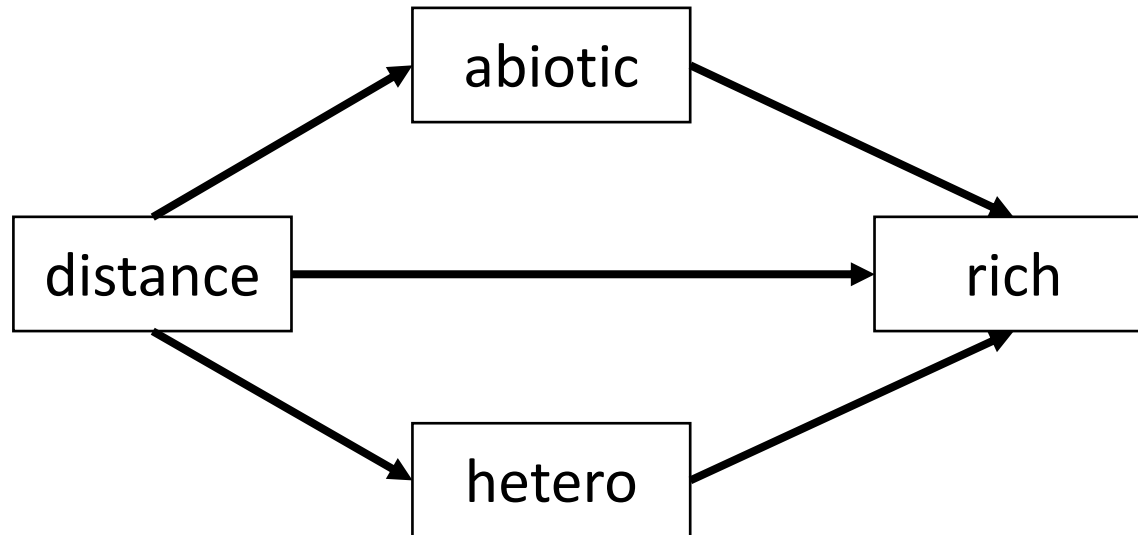
	Independ.Claim	Estimate	Std.Error	DF	Crit.Value	P.Value
1 hetero ~ abiotic + ...		0.002229248	0.001676649	87	1.429585	0.1871306

	Fisher.C	df	P.Value
1	3.352	2	0.187

A red arrow points from the P-value 0.1871306 in the first table to the P-value 0.187 in the second table.

## 1.4 piecewiseSEM. Log-likelihood fit



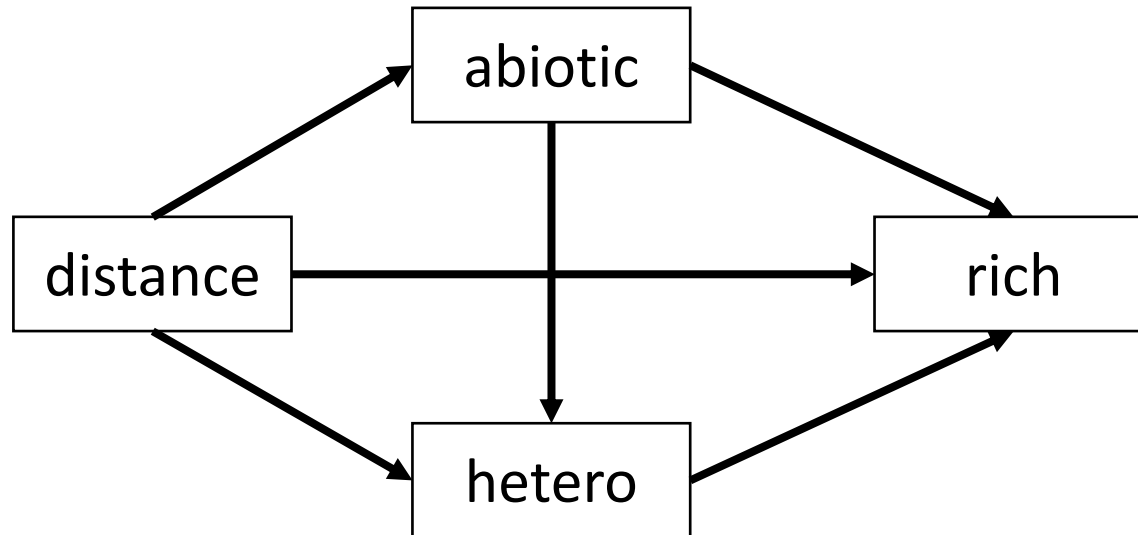
```
# Get log-likelihoods from original model
(M1 <- sapply(keeley.sem2, function(x) ifelse(class(x) == "data.frame",
NA, logLik(x))))
```

			data
-299.9828	73.3324	-342.9844	NA

```
# Sum L-Ls
(M1 <- sum(M1, na.rm = TRUE))
[1] -569.6348
```



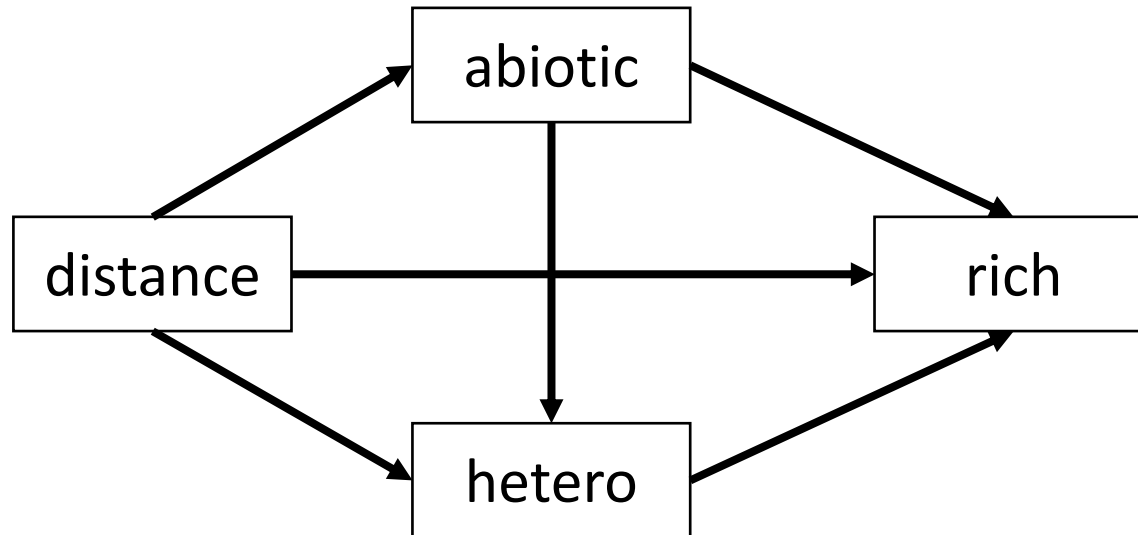
## 1.4 piecewiseSEM. Log-likelihood fit



```
# Fit saturated model (add all missing paths)
Keeley.sem3 <- update(keeley.sem2, hetero ~ abiotic + distance)
```



## 1.4 piecewiseSEM. Log-likelihood fit



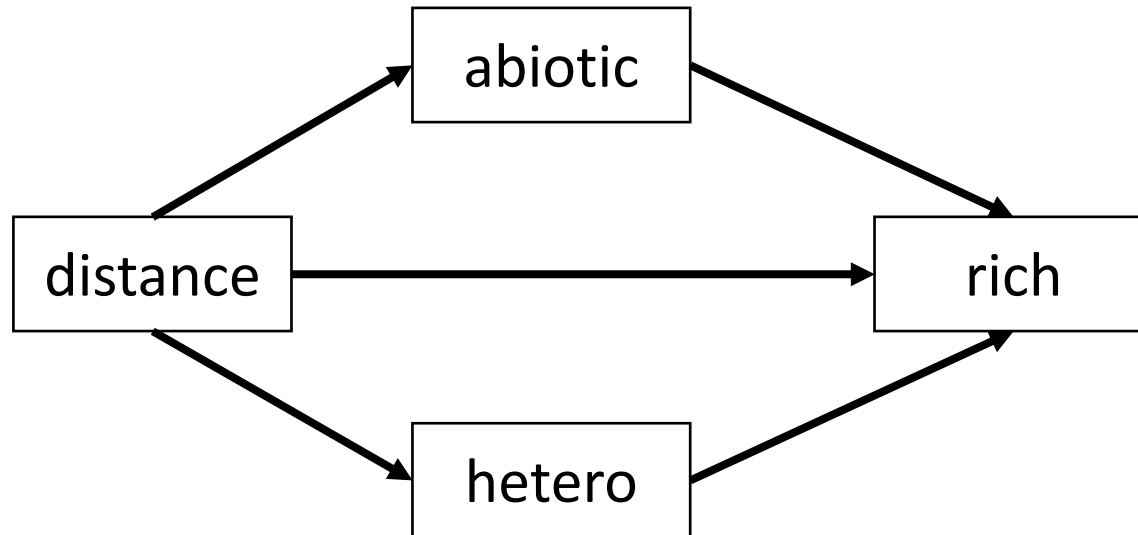
```
# Get log-likelihoods from saturated model  
(M2 <- sapply(keeley.sem3, function(x) ifelse(class(x) == "data.frame",  
NA, logLik(x))))
```

				data
-299.98277	74.23761	-342.98438		NA

```
# Sum L-Ls  
(M2 <- sum(M2, na.rm = TRUE))  
[1] -568.7295
```



## 1.4 piecewiseSEM. Log-likelihood fit



```
# Compute chi-squared statistic  
Chi.sq <- -2*(M1 - M2)
```

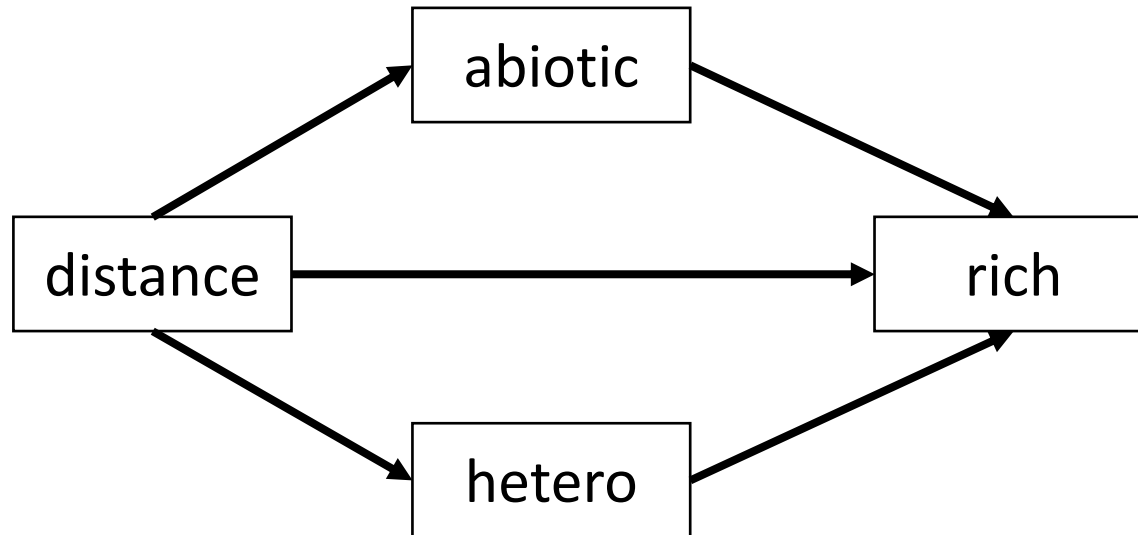
```
# Compare to chi-squared distribution with 1 d.f. (one additional  
estimated parameter in saturated model)  
1 - pchisq(Chi.sq, 1)
```

```
[1] 0.1784574
```





## 1.4 piecewiseSEM. Log-likelihood fit

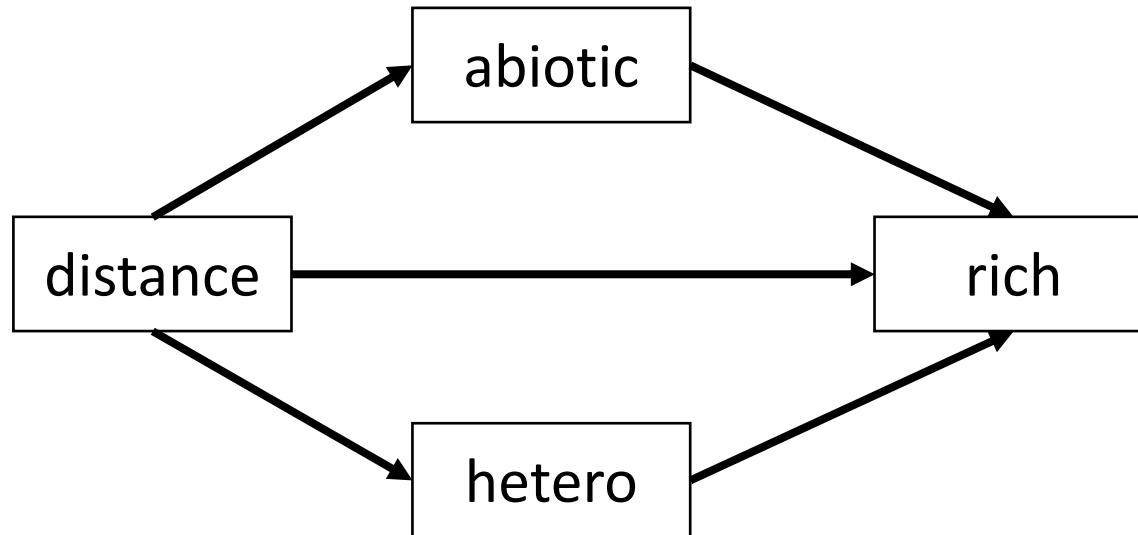


```
# Auto-magic calculation!  
LLchisq(keeley.sem2)
```

	Chisq	df	P.value
1	1.81	1	0.178



## 1.4 piecewiseSEM. Log-likelihood fit



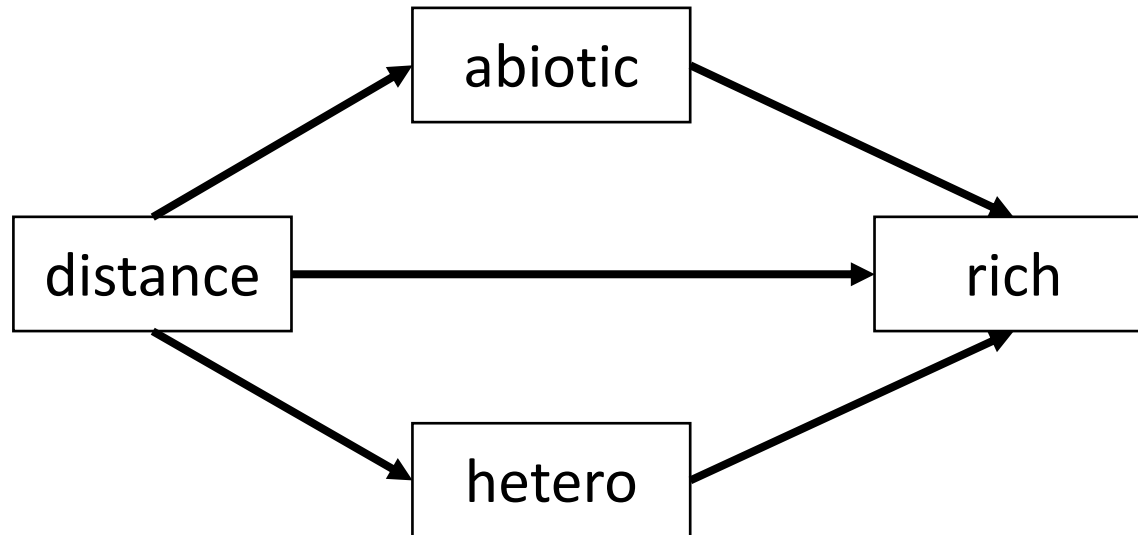
```
# Same P-value as from lavaan (chi-squared value too!)  
model <- '  
abiotic ~ distance  
hetero ~ distance  
rich ~ abiotic + hetero + distance  
'
```

```
lavaan::lavInspect(lavaan::sem(model, keeley), "fit")["pvalue"]
```

```
pvalue  
0.1784574
```



## 1.4 piecewiseSEM. Get coefficients

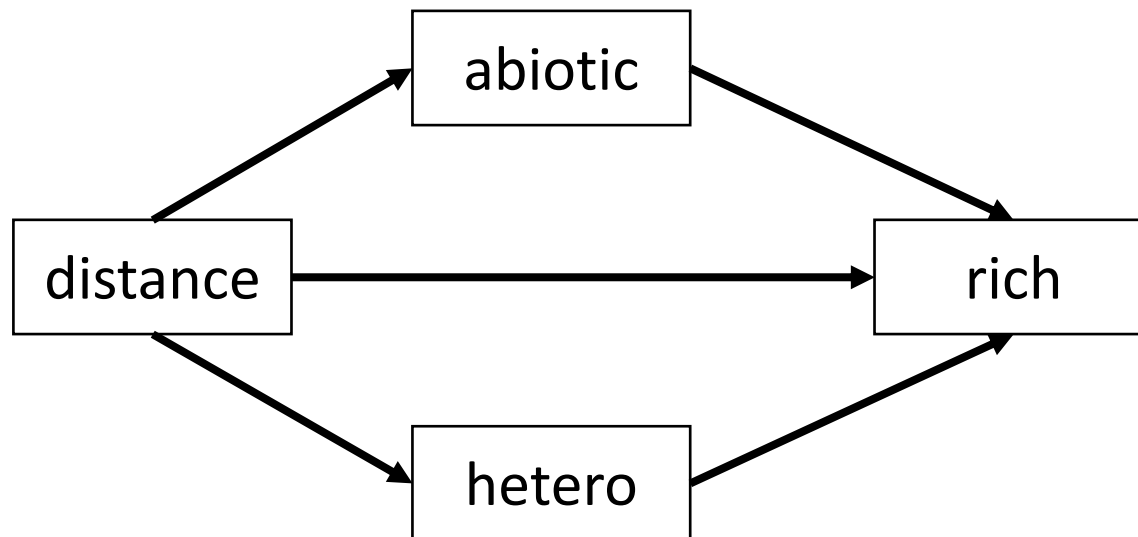


```
# Get coefficients  
coefs(keeley.sem2)
```

	Response	Predictor	Estimate	Std.Error	DF	Crit.Value	P.Value	Std.Estimate	
1	abiotic	distance	0.3998	0.0823	88	1.8562	0.0000	0.4597	***
2	hetero	distance	0.0045	0.0013	88	3.4593	0.0008	0.3460	***
3	rich	abiotic	0.5233	0.1756	86	2.9793	0.0038	0.2660	**
4	rich	hetero	33.4010	11.1187	86	3.0040	0.0035	0.2539	**
5	rich	distance	0.6404	0.1565	86	1.0933	0.0001	0.3743	***



## 1.4 piecewiseSEM. Get coefficients

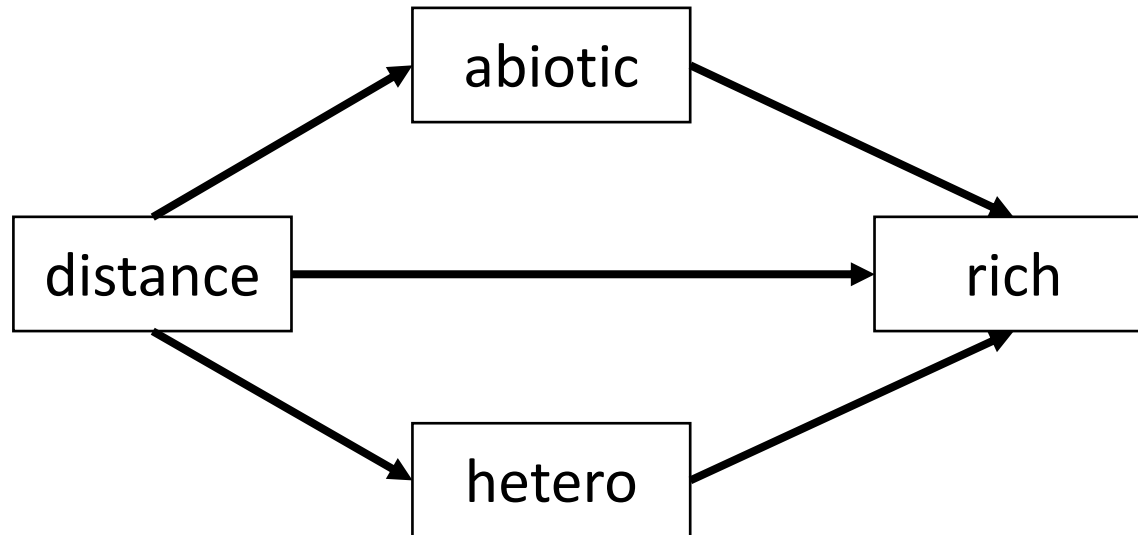


```
# Return intercepts as well  
coefs(keeley.sem2, intercepts = T)
```

	Response	Predictor	Estimate	Std.Error	DF	Crit.Value	P.Value	Std.Estimate	
1	abiotic	(Intercept)	29.5537	4.1176	88	7.1774	0.0000	0.0000	***
2	abiotic	distance	0.3998	0.0823	88	4.8562	0.0000	0.4597	***
3	hetero	(Intercept)	0.4618	0.0650	88	7.0997	0.0000	0.0000	***
4	hetero	distance	0.0045	0.0013	88	3.4593	0.0008	0.3460	***
5	rich	(Intercept)	-30.8880	9.5340	86	-3.2398	0.0017	0.0000	**
6	rich	abiotic	0.5233	0.1756	86	2.9793	0.0038	0.2660	**
7	rich	hetero	33.4010	11.1187	86	3.0040	0.0035	0.2539	**
8	rich	distance	0.6404	0.1565	86	4.0933	0.0001	0.374	***



## 1.4 piecewiseSEM. Get coefficients



```
# Get R-squared  
rsquared(keeley.sem2)
```

	Response	family	link	method	R.squared
1	abiotic	gaussian	identity	NA	0.2113455
2	hetero	gaussian	identity	NA	0.1197074
3	rich	gaussian	identity	NA	0.4700472

# 1.4 piecewiseSEM. Summary

```
# Get all summary information
summary(keeley.sem2)
```

Structural Equation Model of keeley.sem2

```
Call:
  abiotic ~ distance
  hetero ~ distance
  rich ~ abiotic + hetero + distance
```

```
      AIC
1161.270
```

---

Tests of directed separation:

	Independ.Claim	Test.Type	DF	Crit.Value	P.Value
hetero ~ abiotic + ...	coef	87	1.3296	0.1871	

--

Global goodness-of-fit:

Chi-Squared = 1.81 with P-value = 0.178 and on 1 degrees of freedom  
Fisher's C = 3.352 with P-value = 0.187 and on 2 degrees of freedom



# 1.4 piecewiseSEM. Summary

---

Coefficients:

Response	Predictor	Estimate	Std.Error	DF	Crit.Value	P.Value	Std.Estimate	
abiotic	distance	0.3998	0.0823	88	4.8562	0.0000	0.4597	***
hetero	distance	0.0045	0.0013	88	3.4593	0.0008	0.3460	***
rich	abiotic	0.5233	0.1756	86	2.9793	0.0038	0.2660	**
rich	hetero	33.4010	11.1187	86	3.0040	0.0035	0.2539	**
rich	distance	0.6404	0.1565	86	4.0933	0.0001	0.3743	***

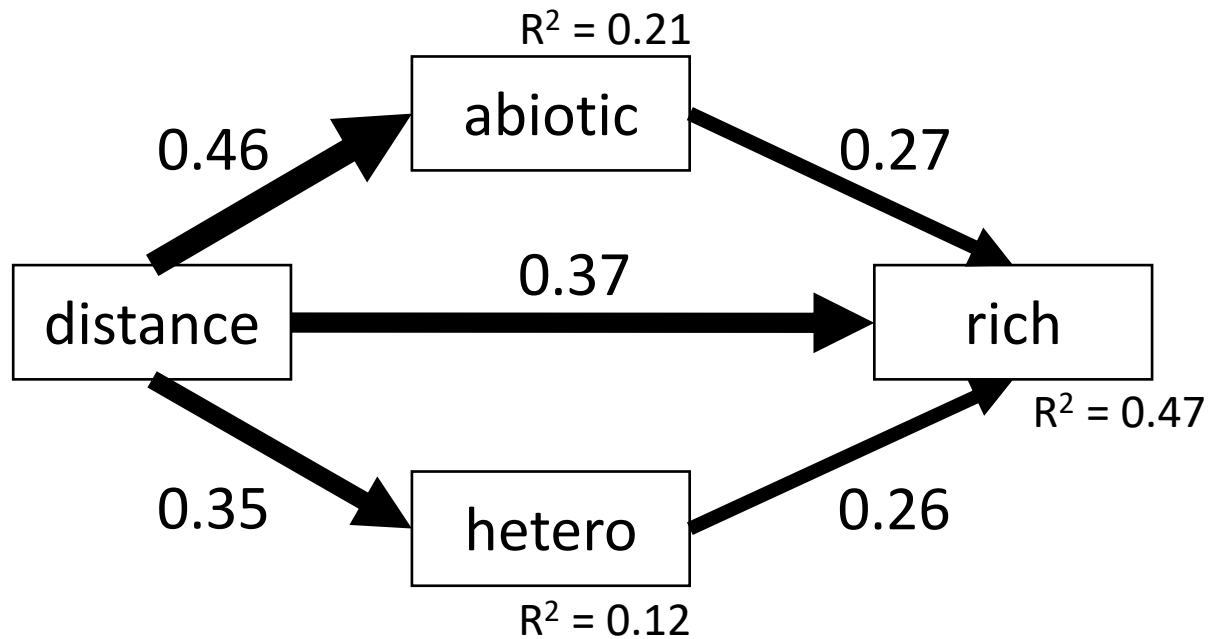
Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05

---

Individual R-squared:

Response	method	R.squared
abiotic	none	0.21
hetero	none	0.12
rich	none	0.47

## 1.4 piecewiseSEM. Summary



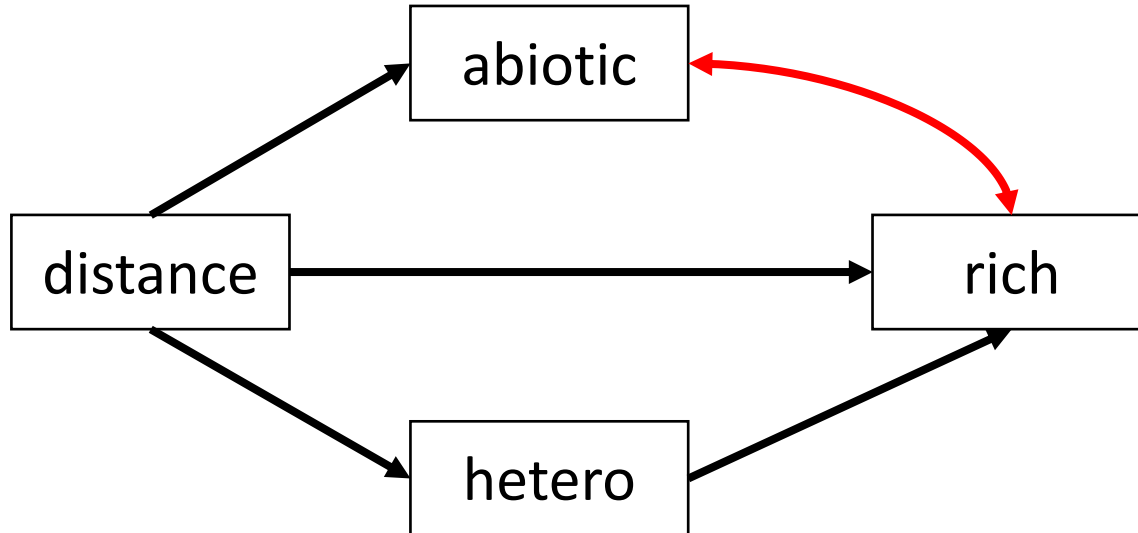


## 1.4 piecewiseSEM. Summary



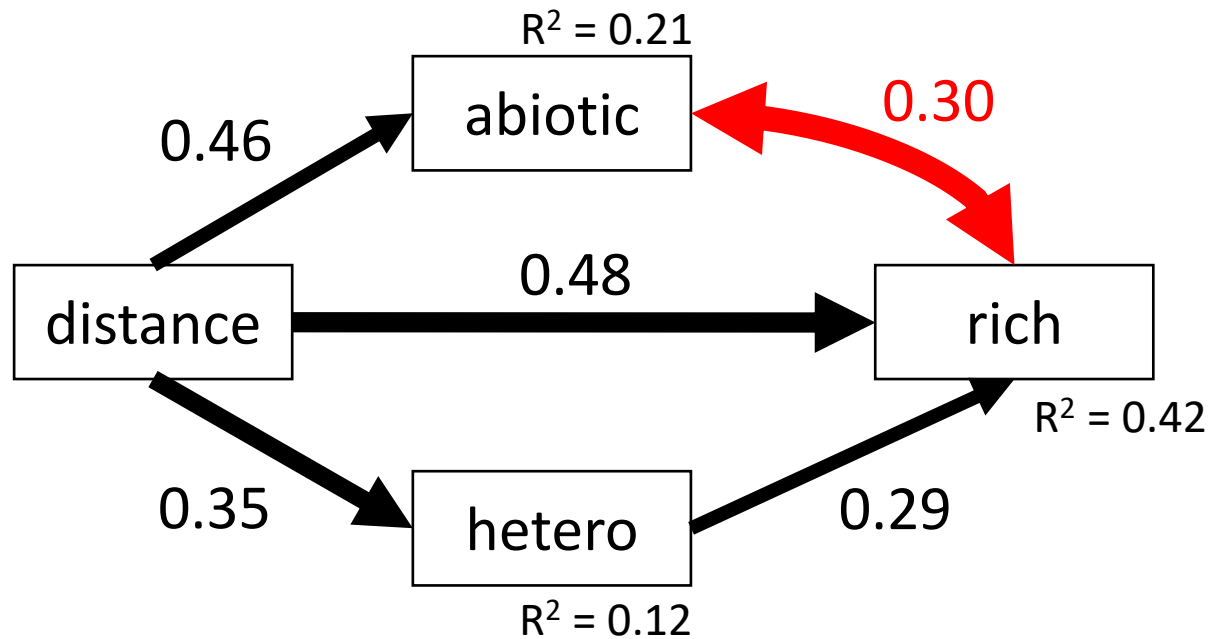
```
# Use built-in plotting function based on `diagrammer`  
plot(keeley.sem2)
```

## 1.4 piecewiseSEM. Correlated errors

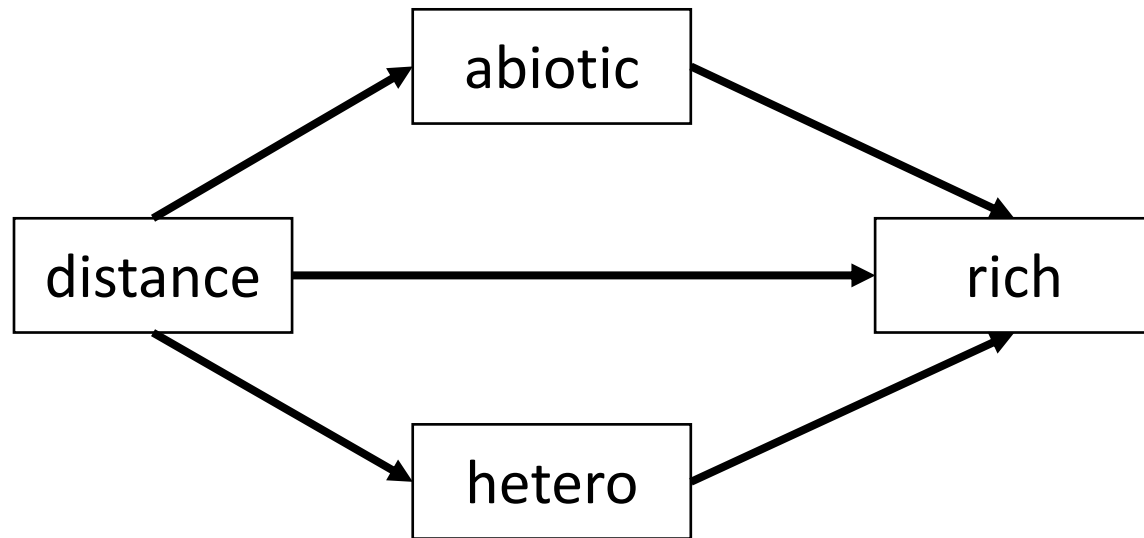


```
keeley.sem3 <- psem(  
  lm(abiotic ~ distance, data = keeley),  
  lm(hetero ~ distance, data = keeley),  
  lm(rich ~ distance + hetero, data = keeley),  
  rich %~~% abiotic # same syntax as lavaan  
)  
  
summary(keeley.sem3)
```

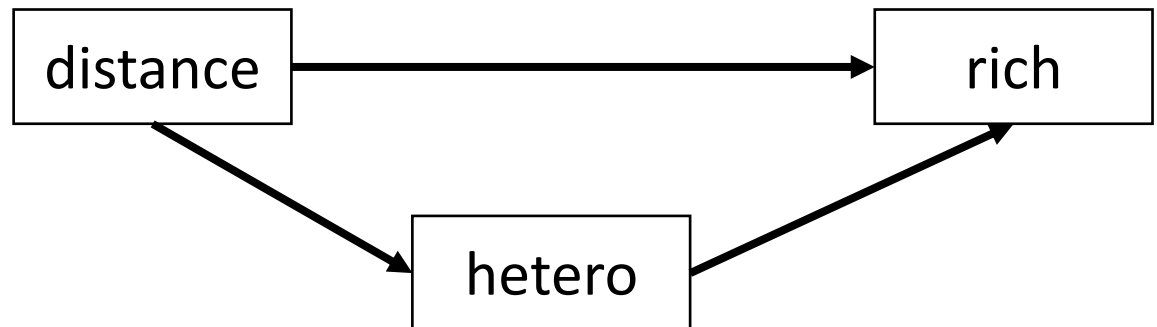
## 1.4 piecewiseSEM. Correlated errors



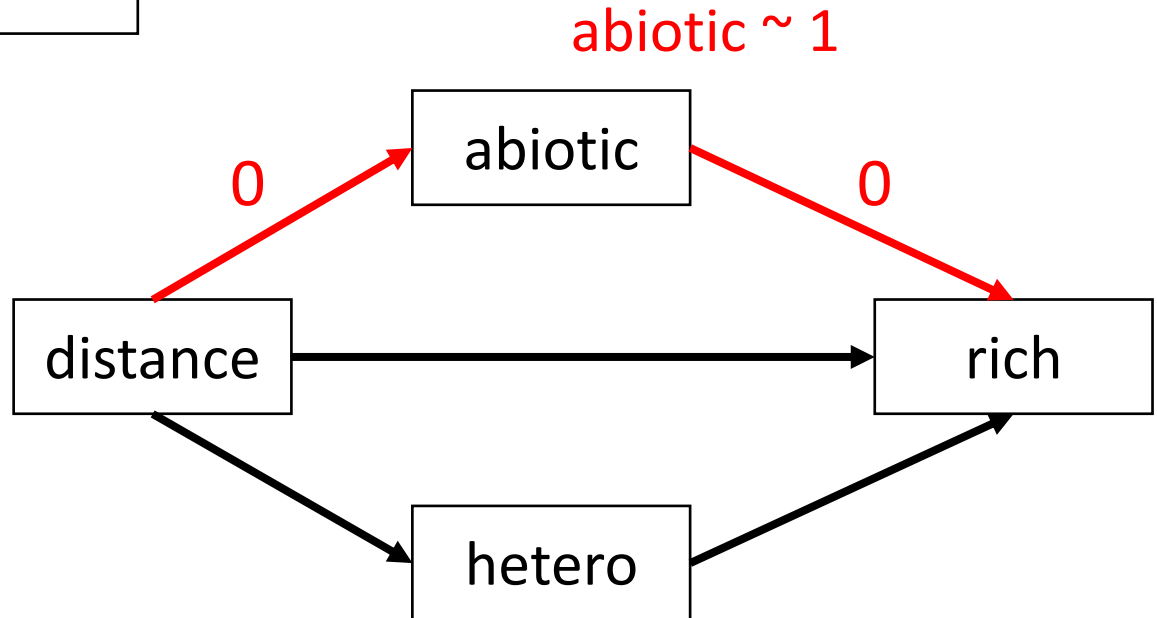
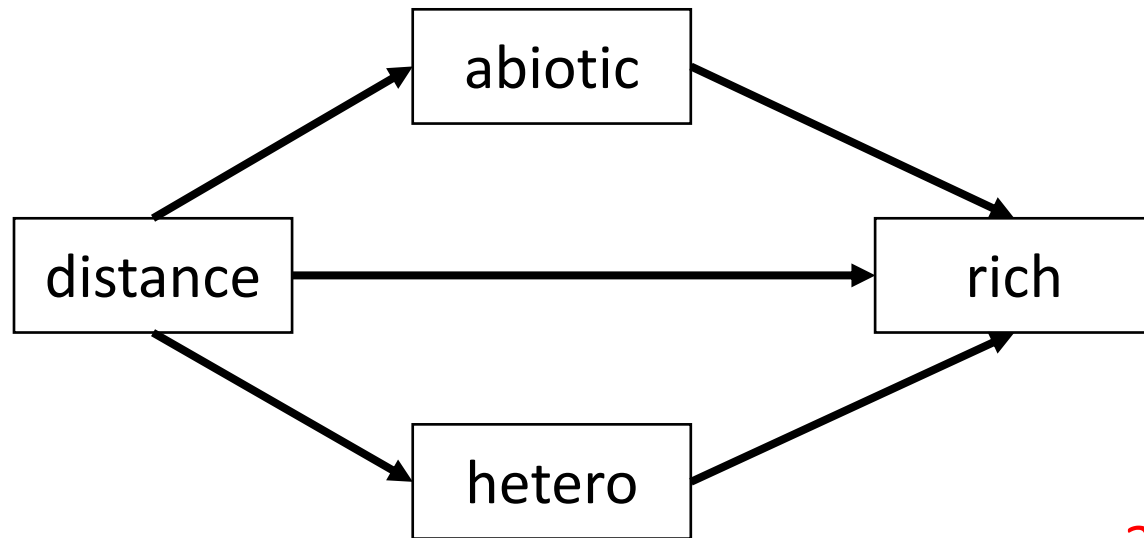
## 1.4 piecewiseSEM. AIC comparisons



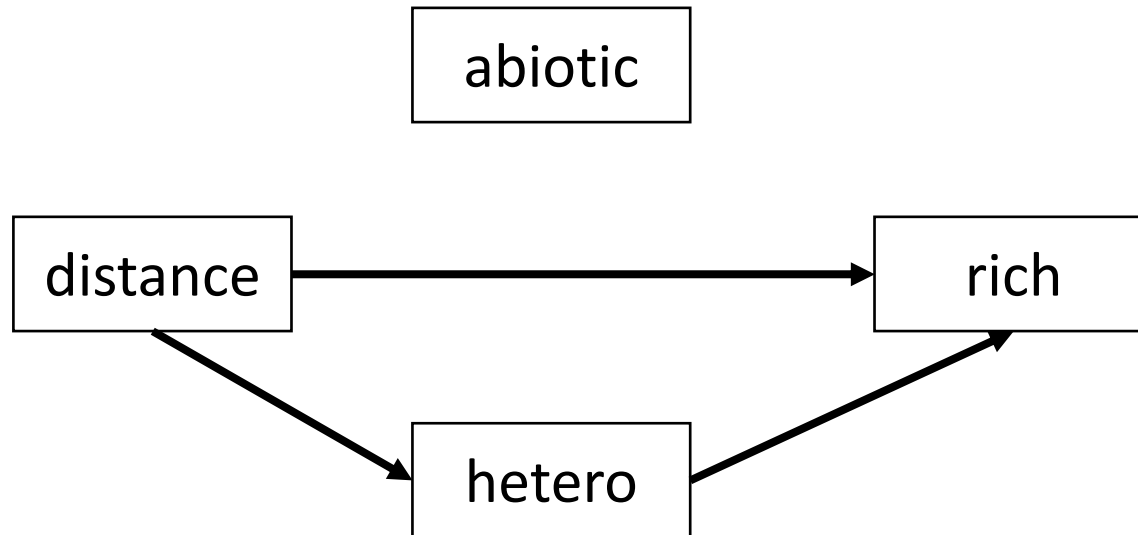
abiotic



## 1.4 piecewiseSEM. AIC comparisons



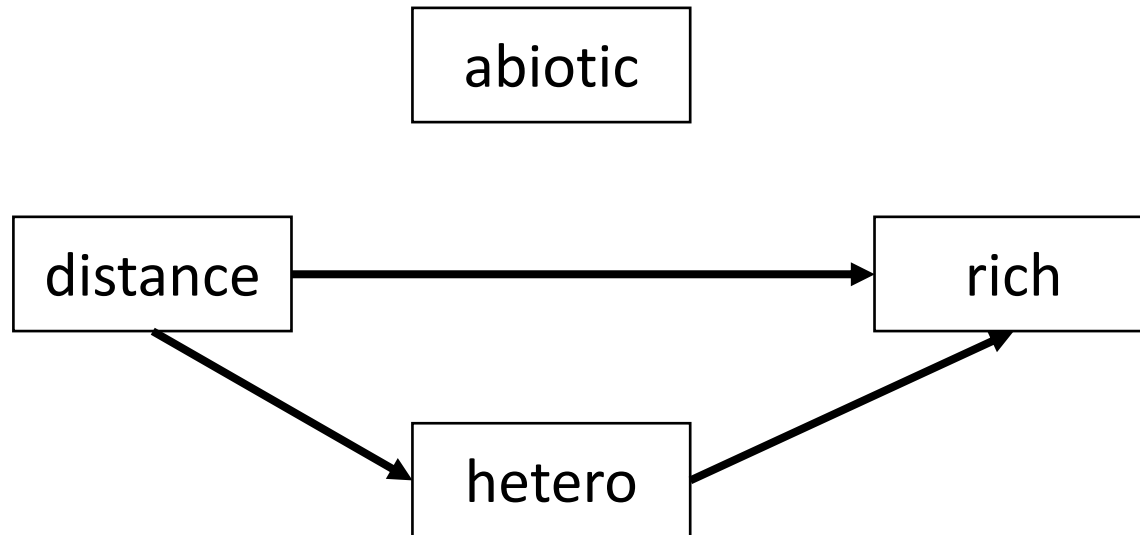
## 1.4 piecewiseSEM. Fit new model



```
# Fit alternate model
keeley.sem4 <- psem(
  lm(hetero ~ distance, data = keeley),
  lm(rich ~ distance + hetero, data = keeley),
  lm(abiotic ~ 1, data = keeley)
)
```



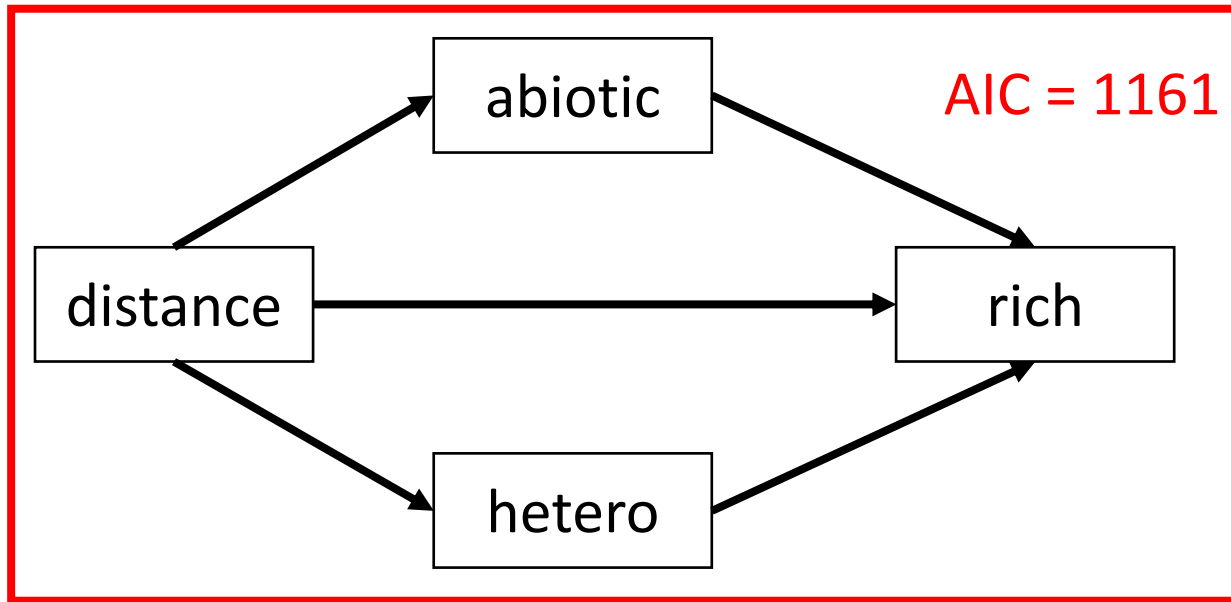
## 1.4 piecewiseSEM. Fit new model



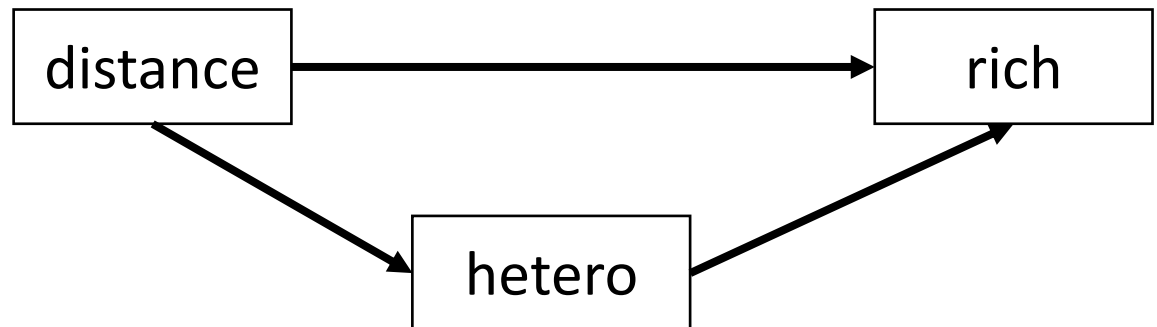
```
# Compare the two models using AIC  
AIC(keeley.sem2, keeley.sem4)
```

	df	AIC
x	11	1161.270
y	9	1187.479

## 1.4 piecewiseSEM. AIC comparisons

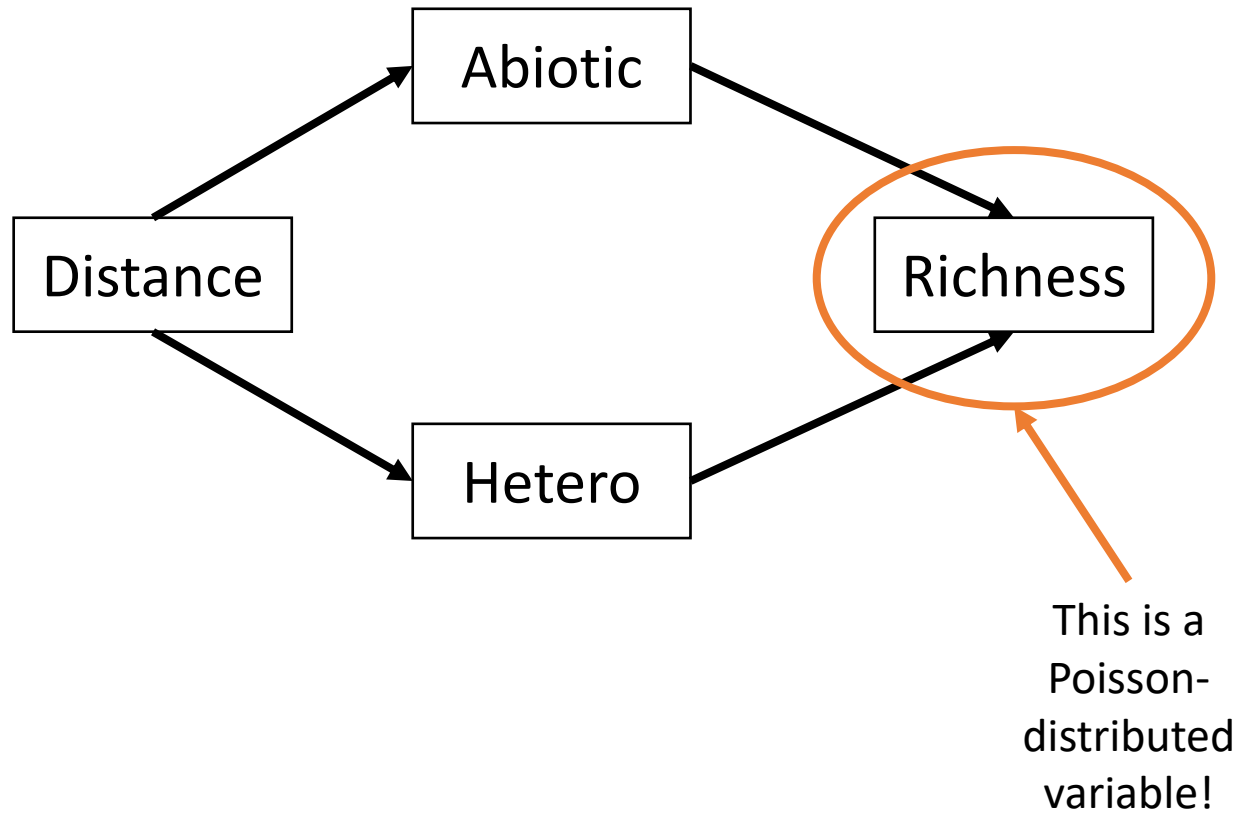


AIC = 1187





## 1.4 SEM Examples. Refit Keeley using GLM



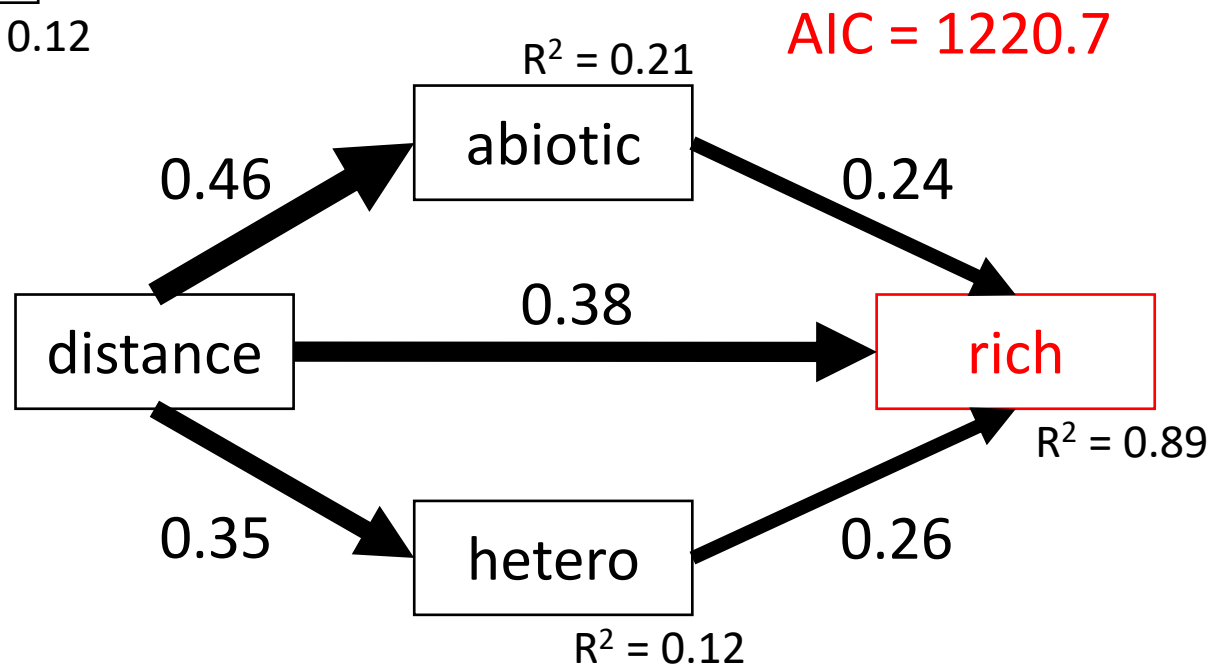
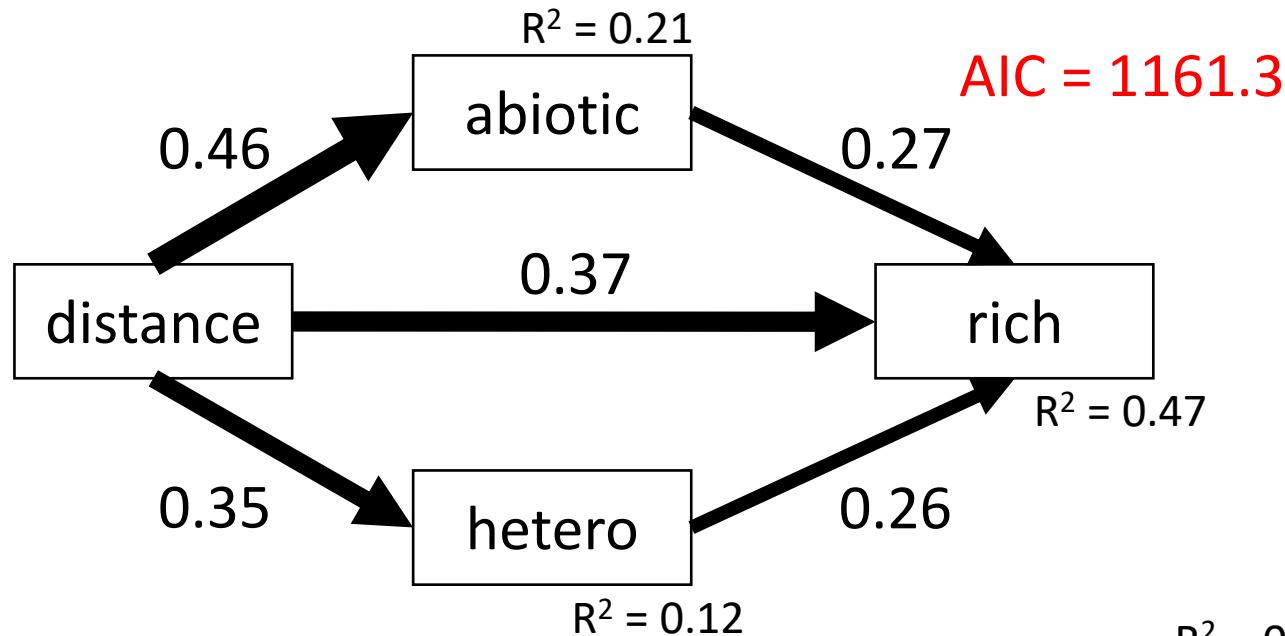
# 1.4 SEM Examples. Refit Keeley using GLM

```
# Re-run Keeley with GLM for richness
keeley.glm.sem <- psem(
  lm(abiotic ~ distance, data = keeley),
  lm(hetero ~ distance, data = keeley),
  glm(rich ~ abiotic + hetero + distance, family = "poisson", data =
keeley),
  keeley
)

summary(keeley.glm.sem)
```



# 1.4 SEM Examples. Refit Keeley using GLM

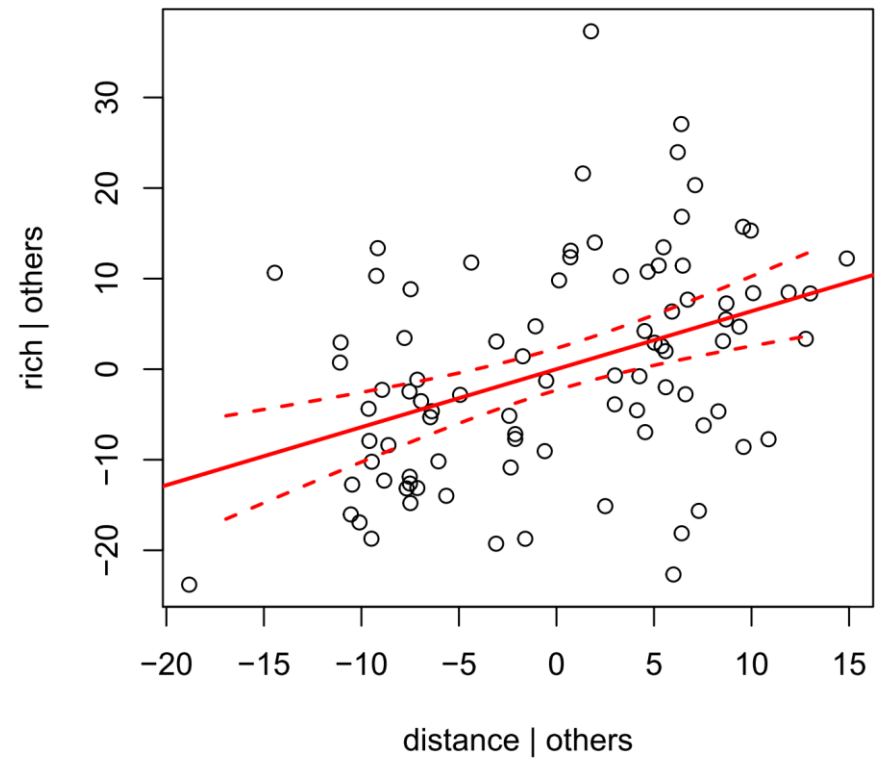
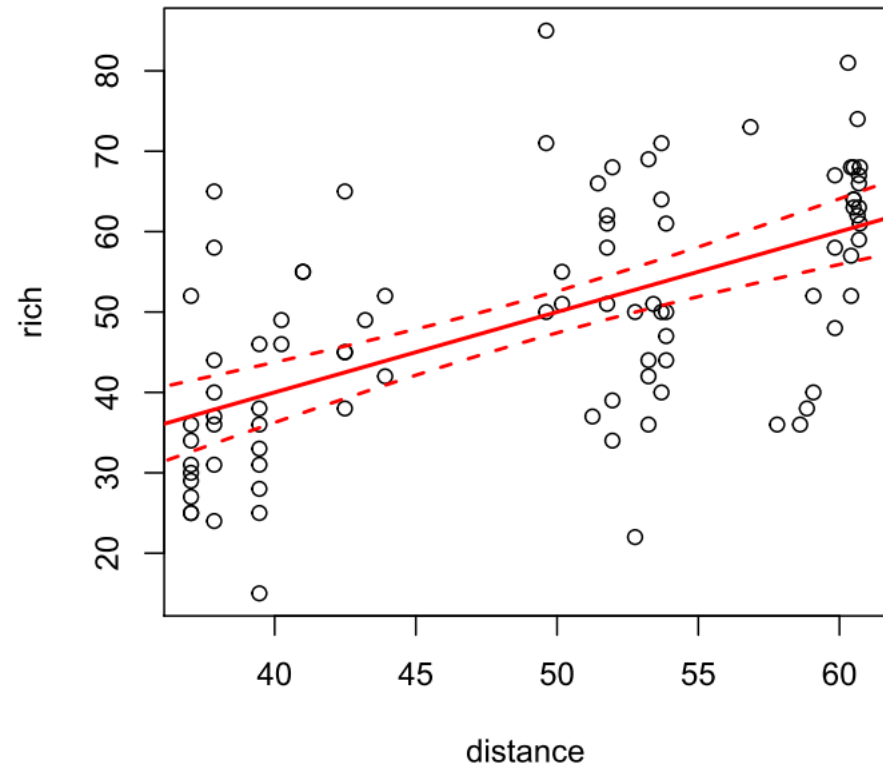


## 1.4 piecewiseSEM. Partial regression coefficient

Isolate the independent effect of distance on richness:

1. Regress abiotic and hetero against richness (removing distance)
2. Regress distance against abiotic and hetero (remove rich)
3. Regression residuals of 1 against 2 (having removed effects of abiotic and hetero from both)

# 1.4 piecewiseSEM. Partial regression coefficient



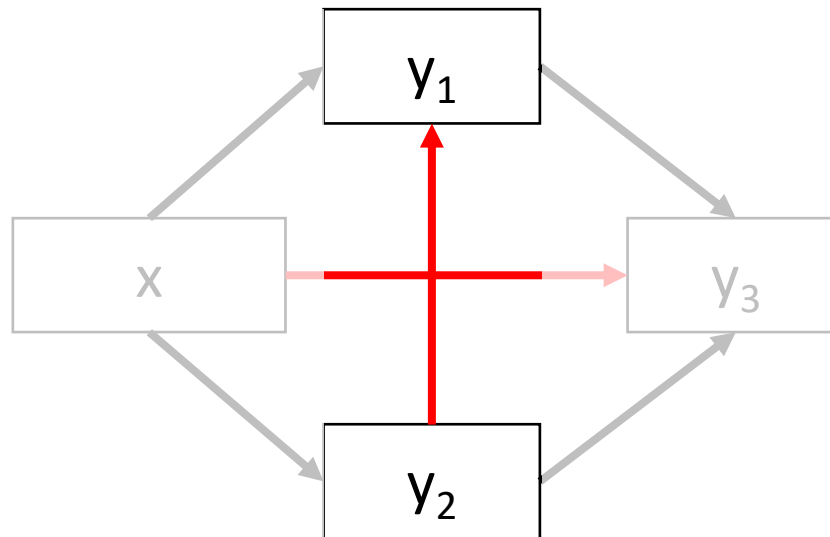
## 1.4 piecewiseSEM. Partial regression coefficient

- Useful for displaying trends, particularly with complex models where bivariate correlations are messy
- Can be used for any multiple regression (single model or list)
- Not applicable to simple regression ( $Y \sim X$ ) for obvious reasons
- Well implemented in *emmeans* package

1.5 A Warning...

## 1.5 Directed Separation. A warning

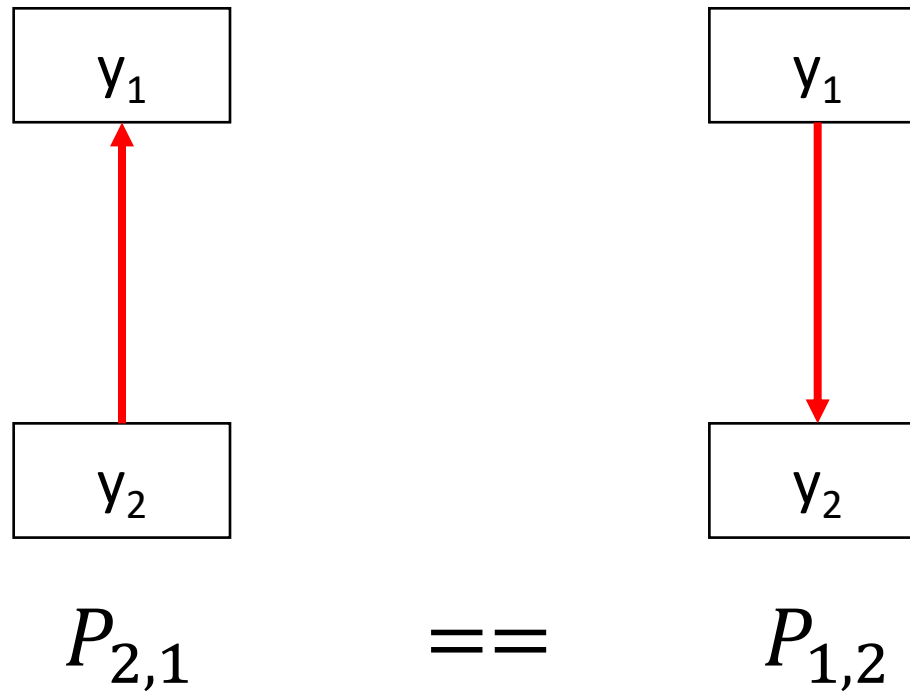
- Intermediate non-normal endogenous variables pose a challenge





## 1.5 Directed Separation. A warning

- If normal, significance values are reciprocal



## 1.5 Directed Separation. A warning

```
set.seed(66)
```

```
data <- data.frame(x = rnorm(100), y1 = rnorm(100), y2 = rpois(100,  
10), y3 = rnorm(100))
```

```
# Show that  $y2 \sim y1$  is the same as  $y2 \sim y1$  for LM
```

```
mody1.y2 <- lm(y1 ~ y2 + x, data)
```

```
mody2.y1 <- lm(y2 ~ y1 + x, data)
```

```
summary(mody1.y2)$coefficients[2, 4]
```

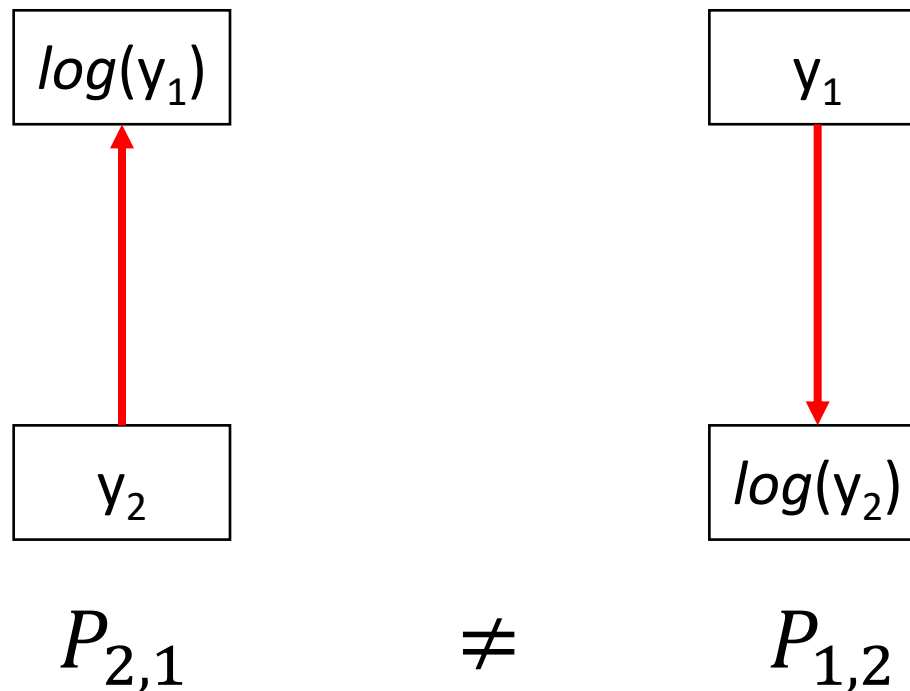
```
[1] 0.7429784
```

```
summary(mody2.y1)$coefficients[2, 4]
```

```
[1] 0.7429784
```

## 1.5 Directed Separation. A warning

- If non-normal, significance values are *not* reciprocal because of transformation via link function



## 1.5 Directed Separation. A warning

```
# Show that  $y_2 \sim y_1$  is not the same as  $y_2 \sim y_1$  for GLM
mody1.y2 <- lm(y1 ~ y2 + x, data)

mody2.y1.glm <- glm(y2 ~ y1 + x, "poisson", data)

summary(mody1.y2)$coefficients[2, 4]

[1] 0.7429784

summary(mody2.y1.glm)$coefficients[2, 4]

[1] 0.8036267
```

## 1.5 Directed Separation. A warning

```
# Same is true for log-likelihoods
```

```
logLik(mody1.y2)
```

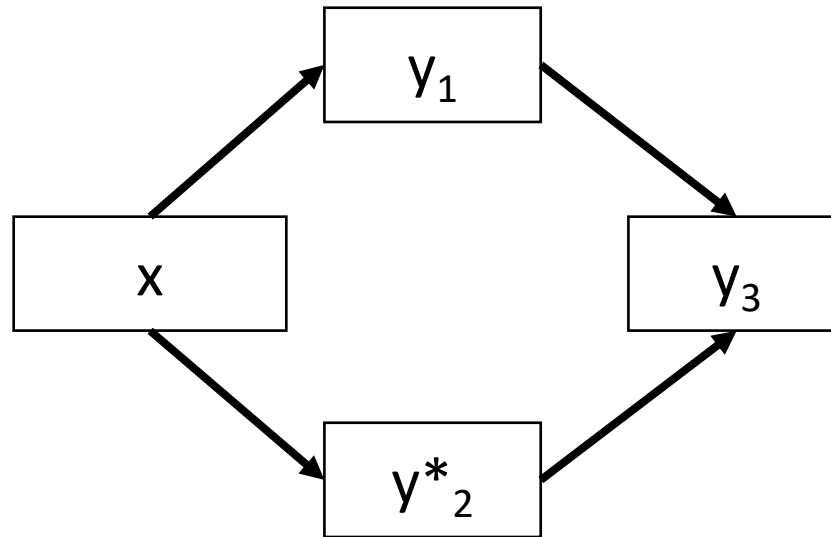
```
'log Lik.' -128.1663 (df=4)
```

```
logLik(mody2.y1.glm)
```

```
'log Lik.' -239.3152 (df=3)
```

```
# Because of differences in ML-fitting function for Gaussian vs.  
Poisson GLM
```

## 1.5 Directed Separation. A warning



```
# Create SEM with GLM
modelList <- psem(
  lm(y1 ~ x, data),
  glm(y2 ~ x, "poisson", data),
  lm(y3 ~ y1 + y2, data),
  data
)
```

## 1.5 Directed Separation. A warning

```
# Run summary  
summary(modelList)
```

Error:

Non-linearities detected in the basis set where P-values are not symmetrical.  
This can bias the outcome of the tests of directed separation.

offending independence claims:

```
y2 <- y1 *OR* y2 -> y1
```

Option 1: Specify directionality using argument 'direction = c()'.  
Option 2: Remove path from the basis set by specifying as a correlated error using '%~~%'.

Option 3: Use argument 'conserve = TRUE' to compute both tests, and return the most conservative P-value.

## 1.5 Directed Separation. A warning

```
# Address conflict using conserve = T
summary(modelList, conserve = T)
```

```
dSep(modelList, conserve = T)
```

	Independ.Claim	Estimate	Std.Error	DF	Crit.Value	P.Value
1	y3 ~ x + ...	-0.01414678	0.09749775	96	-0.1450985	0.8849373
3	y1 ~ y2 + ...	-0.01161551	0.03532167	97	-0.3288495	0.7429784

```
# Check against
summary(mody1.y2)$coefficients[2, 4]
```

```
[1] 0.7429784
```

```
summary(mody2.y1.glm)$coefficients[2, 4]
```

```
[1] 0.8036267
```



## 1.5 Directed Separation. A warning

```
# Address conflict using direction = c()  
dSep(modelList, direction = c("y2 <- y1"))
```

	Independ.Claim	Estimate	Std.Error	DF	Crit.Value	P.Value
1	y3 ~ x + ...	-0.01414678	0.09749775	96	-0.1450985	0.8849373
2	y1 ~ y2 + ...	-0.01161551	0.03532167	97	-0.3288495	0.7429784

```
dSep(modelList, direction = c("y1 <- y2"))
```

	Independ.Claim	Estimate	Std.Error	DF	Crit.Value	P.Value
1	y3 ~ x + ...	-0.01414678	0.09749775	96	-0.1450985	0.8849373
2	y2 ~ y1 + ...	-0.00872099	0.03507248	97	-0.2486562	0.8036267

## 1.5 Directed Separation. A warning

```
# Address conflict using correlated errors  
modelList2 <- update(modelList, y2 %~~% y1)
```

```
dSep(modelList2)
```

	Independ.Claim	Estimate	Std.Error	DF	Crit.Value	P.Value
1	y3 ~ x + ...	-0.01414678	0.09749775	96	-0.1450985	0.8849373

